Physiotherapy rehabilitation following lumbar spinal fusion: a systematic review and meta-analysis of randomised controlled trials

Alison Rushton, Gillian Eveleigh, Emma-Jane Petherick, Nicola Heneghan, Rosalie Bennett, Gill James, Chris Wright

ABSTRACT

Objective: To evaluate the effectiveness of physiotherapy intervention following lumbar spinal fusion.

Design: Systematic review and meta-analysis. 2 independent reviewers searched information sources, assessed studies for inclusion and evaluated risk of bias. Quantitative synthesis using standardised mean differences was conducted on comparable outcomes across trials with similar interventions.

Information sources: Predefined terms were employed to search electronic databases. Additional studies were identified from key journals, reference lists, authors and experts.

Eligibility criteria for included studies: Randomised control trials published in English prior to 30 September 2011 investigating physiotherapy outpatient management of patients (>16 years), following lumbar spinal fusion, with measurements reported on one or more outcome of disability, function and health were included.

Results: 2 Randomised control trials (188 participants) from two countries were included. Both trials included a behavioural and an exercise intervention. 1 trial was evaluated as high risk of bias and one as unclear. 159 participants were incorporated in the meta-analysis. Although evidence from both trials suggested that intervention might reduce back pain short term (6 months) and long term (12 months and 2 years), and a behavioural intervention might be more beneficial than an exercise intervention, the pooled effects (0.72, 95% CI –0.25 to 1.69 at 6 months; 0.52, 95% CI –0.45 to 1.49 at 12 months and 0.75, 95% CI –0.46 to 1.96 at 2 years) did not demonstrate statistically significant effects. There was no evidence that intervention changes pain in the short (6 months) or long term (12 months and 2 years). The wide CI for pooled effects indicated that intervention could be potentially beneficial or harmful. Considerable heterogeneity was evident.

Conclusions: Inconclusive, very low-quality evidence exists for the effectiveness of physiotherapy management following lumbar spinal fusion. Best practice remains unclear. Limited comparability of outcomes and retrieval of only two trials reflect a lack of research in this area that requires urgent consideration.

ARTICLE SUMMARY

Article focus

- Physiotherapy intervention is recommended following lumbar spinal fusion.
- However, the most beneficial intervention and the effectiveness of physiotherapy management are unclear.
- The objective was to evaluate effectiveness of physiotherapy intervention in patients following lumbar spinal fusion on clinically relevant outcomes, short and long term.

Key messages

- Inconclusive, very low-quality evidence exists for the effectiveness of physiotherapy management following lumbar spinal fusion.
- Best practice remains unclear.
- Limited comparability of outcomes and retrieval of only two trials reflect a lack of research in this area that requires urgent consideration.

Strenghts and limitations of this study

- The strengths of this review are its focus to physiotherapy intervention and it being the first in this area; exploring the breadth of potential physiotherapy interventions.
- The key limitation of this review is that differences were evident in the content and nature of interventions, selection of outcome measures and timing of assessment points, contributing to heterogeneity in treatment effects.

INTRODUCTION

Rationale

In the UK National Health Service, surgery is the greatest single component of expenditure for managing low back pain, with increasing numbers of lumbar fusions being performed. More than 4036 operations were...
Physiotherapy rehabilitation following lumbar spinal fusion

performed in 2009/2010, reflecting a 14% increase from 2008/2009. The USA has also seen a well-documented increase in lumbar fusion surgery rates from 1990 to 2001 of 220%, with a corresponding 500% increase in spending for lumbar fusion from 1992 to 2003. The increased rates are partially attributable to advances in technology, including the Food and Drug Administration approval of intervertebral cage implants (1996) and pedicle screws (1998). Although overall lumbar surgical rates in the USA reduced from 2002 to 2007, fusion rates increased from 1.3 to 10.9 per 100 000 patients. In 1992, lumbar fusion accounted for 14% spending on back surgery in the USA, and by 2003, this had increased to 47%. Additional contributions to this increasing problem include more than 200 lumbar fusion revision operations performed annually in the UK, with a re-hospitalisation rate of 13% within 30 days of surgery documented in the USA.

The primary indication for lumbar fusion is pain (back and/or leg pain) from joints with degenerative disease. Lumbar fusion is thought to stabilise the spine and reduce the need for further surgery. A Cochrane review of spinal surgery for lumbar spondylosis due to degenerative causes identified trials of variable quality, with an emphasis on surgical rather than patient outcomes, and little information available regarding occupational or long-term outcomes. The review concluded that there were conflicting results for surgery. These findings were confirmed by Sogaard et al who summarised the literature and also concluded that there are minimal data on the reported success for patient outcome following lumbar fusion. Data from the Swedish National Spine Register reported that 25% patients reported no change or worsened pain following lumbar fusion (back and/or leg pain) and that at 12 months following surgery, 40% of patients reported dissatisfaction regarding the outcome of the surgery.

Re-operation rates have recently become a focus of investigation. Martin et al investigated the cumulative incidence of second operation for degenerative conditions in one USA state, finding that an increased proportion of fusion operations and the technical development of implants did not affect the rate of re-operation. Indeed, surgery in the late 1990s was more likely to be repeated than that in the early 1990s, contributing to a ‘substantial’ likelihood of re-operation.

The existing variability of evidence to evaluate efficacy of lumbar fusion and some evidence of persisting symptoms and dissatisfaction following surgery highlights the necessity for evidence of the effectiveness of post-operative rehabilitation. Evidence for rehabilitation following surgical intervention in low back pain is an area of increasing interest, for example, post-discectomy. There is some debate in the literature regarding timing of early exercise to overload internal fixation. In view of this, Christensen et al commenced rehabilitation after 3 months, although Rohlmann et al found no evidence of compromise of internal fixation through exercises, except perhaps through walking as an exercise.

There is no systematic review investigating effectiveness of rehabilitation in a post-lumbar fusion population. Although physiotherapy intervention is recommended post-lumbar fusion, its effectiveness is unclear, with no evaluation of existing trials through a systematic review. Consequently, current practice and best physiotherapy practice are unclear.

Objective
To investigate the short- and long-term effectiveness of physiotherapy outpatient management following lumbar spinal fusion in terms of disability, function and health in patients aged >16 years.

MATERIALS AND METHODS
A protocol following method guidelines by the Back Review Group of the Cochrane Collaboration and Cochrane handbook informed the conduct of a systematic review, which is reported in line with the PRISMA statement.

Eligibility criteria
Studies
Randomised control trials that evaluated the effectiveness of physiotherapy outpatient management of patients following lumbar spinal fusion.

Participants
Patients who had undergone lumbar spinal fusion, with no complications, aged >16 years.

Interventions
Any physiotherapy outpatient management intervention.

Outcome measures
Measurements reported on one or more outcome of disability, function and health in the short term (approximately 3–6 months post-surgery/intervention) and/or long term (12–24 months).

Information sources
The search employed sensitive topic-based strategies designed for each database (to 30 September 2011):

- The Cochrane Library: Controlled Trials Register, Health Technology Assessment Database, NHS Economic Evaluation Database.
- CINAHL, EMBASE, MEDLINE, PEDro, ZETOC Databases.
- Selected Internet sites and Indexes: Turning Research into Practice, Health Services/Technology Assessment, PubMed.
- National Research Register, Current Controlled Trials website (York).
- Cochrane Back Review Group.
- Hand searches key journals.
Science Citation Index and Social Science Citation Index.


Search
Predefined search terms were used. Box 1 details an example of searches used: the Medline OvidSP search.

Study selection
Two subject experts (GE/E-JP to 15 March 2011 and GE/RB updated to 30 September 2011) searched information sources independently and assessed identified studies for inclusion, facilitated by grading each criterion (table 1) as eligible/not eligible/might be eligible. The full text of a study was reviewed and the study considered potentially relevant when it could not be clearly excluded on the basis of its Title and Abstract following discussion between the two independent reviewers. Full text was obtained for abstracts with insufficient information or in a situation of disagreement. A study was included when both reviewers independently assessed it as satisfying the inclusion criteria from the full text. A third reviewer (AR, methodological and subject expert) mediated in the event of disagreement following discussion.

Risk of bias for each included trial was independently assessed by the same initial reviewers. Consistent with Cochrane, risk of bias and homogeneity of participants, interventions and outcome measures were important considerations informing potential inclusion of trials in meta-analyses, thereby ensuring meaningfulness of findings from a clinical perspective. The third reviewer mediated in situations of disagreement. Cohen’s κ was used to assess agreement between reviewers. All tools and processes were piloted prior to use.

Data collection process
Using a standardised form, two reviewers (AR/CW) extracted the data independently. A third reviewer (NH) independently checked for consistency and clarity.

Data items
Data extracted for each trial included design, participants and indication, interventions and study setting, outcome measures, timing of assessments, and main results. Key outcome measures of interest were predefined as valid tools to measure pain, disability, function, physical impairment, social impact and patient satisfaction, as reflected in the domains from the WHO’s International Classification of Functioning, Disability and Health.

Risk of bias in individual studies
The Cochrane ‘risk of bias’ assessment tool was used to assess the internal validity of each included trial. This approach was developed through empirical research and unlike most quality scales. Each ‘risk of bias’ component was reported independently, in relation to each outcome measure. Assessment by reviewers acknowledged that the component of ‘blinding’ the treating therapist is generally impossible and the Cochrane tool permitted evaluation of the likely influence of lack of binding.

Summary measures
In accordance with the protocol, quantitative synthesis was conducted on comparable key outcomes across trials that had similar interventions (nature of intervention) and timing of assessments (at approximately 6 months, 12 months and 2 years post-surgery). Tools developed to measure the same underlying domain of the International Classification of Functioning were defined as comparable outcomes. Subject and methodological experts (AR/CW/GJ) identified the combinations of trials and outcomes on which to conduct meta-analyses. All results were reported in the context of overall risk of bias for a trial.
Meta-analyses, conducted through RevMan, compared standardised differences in means using DerSimonian—Laird random effects\textsuperscript{28 29} as the principal analyses to allow for systematic differences in effects estimated across the included trials.\textsuperscript{19 29} Ninety-five per cent confidence intervals were reported for summary statistics. For comparisons across trials that reported different measurement tools for the same outcome\textsuperscript{19} or a mixture of final value scores and change from baseline scores, standardised mean differences were selected.\textsuperscript{30}

**Planned methods of analysis**

All authors were contacted to request either raw data or additional summary statistics to those reported. No raw data were supplied, and analyses were therefore conducted on the reported final summary statistics. Standard deviations were estimated from reported CIs or percentiles, as necessary.\textsuperscript{30} Change scores were used for studies when no other data were forthcoming, in-line with the use of random effects as primary analyses.\textsuperscript{29} Heterogeneity in treatment effects was considered by computation of $I^2$. An analysis of the quality of the interventions was undertaken as the basis for interpretation of heterogeneity.\textsuperscript{19 31}

**Risk of bias across trials**

Retrieval of too few trials reporting comparable outcome measures precluded visual assessment of potential publication bias using Funnel plots.\textsuperscript{19} Consensus regarding the overall potential risk of bias was facilitated through tabulation of the summary assessment for risk of bias.

**Additional analyses**

With only two trials included in the review, there was a lack of information on which to conduct post-hoc supportive analyses beyond descriptive analysis.

**RESULTS**

**Study selection**

In total, four articles\textsuperscript{8 13 32 33} and two trials\textsuperscript{13 33} from two countries were included. For one trial,\textsuperscript{13} two further articles retrieved presented additional data to the original trial and were considered as part of the main trial.\textsuperscript{8 32} All but one of the retrieved trials were published in English. No relevant unpublished studies were found. Figure 1 presents the numbers of studies at each stage of selection. Complete inter-reviewer agreement was achieved on study inclusion following discussion.

**Study characteristics**

Descriptive data for the two included trials are summarised in Table 2.

**Methods**

Abbott et al\textsuperscript{33} randomised participants across two groups and compared exercise therapy with psychomotor therapy. Christensen et al\textsuperscript{8} randomised participants across three groups and compared a video of home exercises, a back-cape and a training intervention. Both trials, therefore, included a behavioural and an exercise intervention. Duration of interventions ranged from 1 day to 9 weeks, starting between 1 day and 3 months post-surgery. Abbott et al\textsuperscript{33} therefore, commenced interventions during the patient’s inpatient stay. The
number of assessments varied from 3 to 4, occurring 3 months to 2/3 years post-surgery.

Participants
Although the population for trial inclusion was broad and encompassed all patients following lumbar spinal fusion, both trials included patients of primarily a degenerative cause. The two trials randomised a total of 188 participants. Age varied from 18 to 65 years. One hundred and fifty-nine participants were included in the meta-analyses (table 2) that omitted one of the three groups from Christensen et al.13

Interventions
Both trials were conducted at single centres. Settings for interventions ranged from home to inpatient wards and to outpatient physiotherapy clinics/sections of the hospital. One trial13 included at least one intervention that was a group intervention (back-cafe), although it was unclear whether their exercise training intervention was also a group intervention. All other interventions were individual 1:1 interventions. Interventions could be grouped according to whether they were focused on exercise or a behavioural approach. Exercise interventions included exercise therapy commenced in an inpatient ward and progressed prior to discharge to provide a programme of home exercises33 and a 1.5 h training session twice a week.13 Both exercise interventions included components in the hospital environment (as an inpatient or outpatient) and one included home exercises.33 Behavioural interventions included psycho-motor therapy using cognitive behavioural principles in addition to exercise33 and a back-cafe using physiotherapist and group support to continue exercises. Timing of interventions ranged from 1 day to 9 weeks, starting between 1 day and 3 months post-surgery.

Primary outcomes
One trial specified a primary outcome: the Oswestry Disability Index V.2.0.33

Secondary and additional outcomes
Both trials reported some assessment of back pain, psychological functioning and measures of occupational outcome. No other comparable outcome measures were used across the two trials. One aspect of the psychosocial outcomes, the influence of back pain on future life, was reported by both trials. However the quality of the outcome was unclear in one trial.13 Other outcomes included the following: Visual Analogue Scale back pain, subcomponents of the Low Back Pain Rating Scale (back and leg pain; function and psychological capacity), European Quality of Life Questionnaire EQ-5D, SF-36 mental health subscale, Self-efficacy Scale, Back Beliefs...
<table>
<thead>
<tr>
<th>Trial</th>
<th>Design</th>
<th>Participants and indication</th>
<th>Intervention and setting</th>
<th>Outcome measures</th>
<th>Main results</th>
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</tr>
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<tbody>
<tr>
<td>Abbott et al (2010)</td>
<td>RCT</td>
<td>Patients aged 18–65 years, presenting with back pain/sciatica &gt;12 months without success for</td>
<td>Intervention: A: Exercise therapy, inpatient respiratory and circulatory exercises, training in transfers, walking and other activities of daily living. Prior to discharge, 20 min instruction in home exercises including dynamic endurance exercises for back, abdominal and leg muscles, stretches, and cardiovascular exercises. Progression of intensity and quantity by patient’s perceived level of pain. Home programme continued from 0 to 12 weeks. Activity restriction from contact sports, running, heavy lifting, outer-range lumbar movements for 6 months post-operation.</td>
<td>Short term: ODI V.2.0 (primary outcome)</td>
<td>Statistically significant improvement for group B compared with group A in: ODI, SES, BBQ and TSK at 3, 6 and 12 months, and 2–3 years; VAS at 3 and 6 months; EQ-5D at 12 months; CSQ-catastrophising subscale in group B at 6 months and 2–3 years; CSQ coping to control pain and ability to decrease pain subscales at 3, 6 and 12 months. More patients employed 2–3 years after surgery. And fewer patients had sickness leave duration &gt;6 months after surgery. Authors did respond to request for data, but no additional data available.</td>
<td>Primary outcome measure specified. No primary end point specified. A priori power calculation conducted on ODI (n=0.05; power =90%, MCID =10). Loss to follow-up: Dropout: At 3 months (6%): A: n=3 withdrew B: n=3 treatment protocol violation No dropouts at 6-month follow-up. At 12 months (7%): A: n=2 (n=1 more pain, n=1 re-operation) At 2–3 years (19%): A: n=4 (n=1 not stated, n=3 pain free) B: n=8 (n=2 deceased, n=3 pain free, n=1 more pain, n=2 re-operation) No exclusions Management of losses: missing values imputed taking into account level of pain, when dropout was not associated with pain. Co-interventions not explored. ITT analyses performed.</td>
</tr>
<tr>
<td>Sweden</td>
<td>Two groups: A: Exercise therapy, B: Psychomotor therapy</td>
<td>Recruitment strategy: patients selected for lumbar spinal fusion by spinal surgeons at one University Hospital orthopaedic clinic. Recruitment over 2-year period 2005–2007.</td>
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Physiotherapy rehabilitation following lumbar spinal fusion

### Table 2 continued

<table>
<thead>
<tr>
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<td>A: Video A: Video</td>
<td>View video demonstration of rehabilitation exercises, followed by one oral instruction by a physiotherapist. Exercises designed to provide dynamic muscular training to enhance endurance back, abdominal, and leg muscles, stretching, warm up, and restrictions to activity (no contact sports, training on machines at a fitness centre, jogging). Video and written instructions provided.</td>
<td>Patients post-lumbar spondylosis or stenosis previously, with severe leg pain, who have chronic low back pain caused localised lumbar or lumbosacral segmental instability caused by isthmic spondylolisthesis I/II, primary degeneration (no previous surgery), secondary degeneration after decompression surgery; n = 20 initially included in each group to ensure validity. n = 5 lost to follow-up to give n = 85.</td>
<td>Work LBPRS: Low Back Pain Rating Scale</td>
<td>Statistically significant differences between groups for climbing stairs without pain at baseline.</td>
<td>No management included in each group.</td>
</tr>
<tr>
<td>B: Back-cafe</td>
<td>As A with additional back-cafe meetings with a physiotherapist and other patients to exchange experiences, doubts, reassurances, tips, and psychological support. Involved tea/coffee, 10 min small talk, duration 1.5 h, three over period of 8 weeks.</td>
<td>Patients post-lumbar spondylosis or stenosis previously, with severe leg pain, who have chronic low back pain caused localised lumbar or lumbosacral segmental instability caused by isthmic spondylolisthesis I/II, primary degeneration (no previous surgery), secondary degeneration after decompression surgery; n = 20 initially included in each group to ensure validity. n = 5 lost to follow-up to give n = 85.</td>
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<td>C: Training</td>
<td>Warm up of 15 min, condition training to improve heart lung function, dynamic muscular endurance training of the back, abdomen, and legs, and stretching exercises. Lasted 1.5 h, twice weekly over 8 weeks, progressed repetitions individually. Unclear if a group intervention.</td>
<td>Patients post-lumbar spondylosis or stenosis previously, with severe leg pain, who have chronic low back pain caused localised lumbar or lumbosacral segmental instability caused by isthmic spondylolisthesis I/II, primary degeneration (no previous surgery), secondary degeneration after decompression surgery; n = 20 initially included in each group to ensure validity. n = 5 lost to follow-up to give n = 85.</td>
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**Outcome measures**
- Work LBPRS: Low Back Pain Rating Scale
- LBPRS: Function and psychological capacity
- LBPRS: Back and leg pain
- LBPRS: Work status
- Back-related healthcare

**Main results**
- Statistically significant improvements for:
  - leg pain in groups A and B compared with group C at 24 months.
  - Components of function for group B compared with groups A and C at 24 months.
  - Sick leave for group B compared with groups A and C at 12 and 24 months.

**Comments**
- Short term: Work LBPRS: back and leg pain
- Long term: LBPRS: function and psychological capacity

**Setting**
Previously, with severe leg pain, who have chronic low back pain caused localised lumbar or lumbosacral segmental instability caused by isthmic spondylolisthesis I/II, primary degeneration (no previous surgery), secondary degeneration after decompression surgery; n = 20 initially included in each group to ensure validity. n = 5 lost to follow-up to give n = 85.

### Table 3

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</tbody>
</table>
Risk of bias within studies

Good inter-reviewer agreement was achieved on risk of bias (Cohen’s $k = 0.613$, 95% CI 0.359 to 0.868).²¹ Of the two included trials, one was evaluated overall as high risk of bias¹³ and one as unclear³³ (table 3). Risk of bias was, therefore, considered in conjunction with other indicators of study differences (comparability of interventions, outcome measures and timings of assessments) to determine any appropriate quantitative synthesis of the trials.¹⁹ Interestingly, in the subsequent reporting of the Christensen trial,⁸ ³² the risk of bias was improved, although, overall, it remained high (table 5). This suggests that poor reporting contributed to the rating of high risk of bias for multiple issues in the original trial report.

Risk of bias across studies

Christensen et al²³ had one high-risk component owing to poor reporting affording a lack of clarity across all components, no primary outcome being pre-specified, no primary end point being pre-specified and no intention-to-treat analysis reported. Both trials reported losses to follow-up. However, in both trials, losses were <20% and evaluated as acceptable.³⁴ Interpretation of results could be affected by the high proportion of information from one trial identified as high risk of bias.²³

Results of individual trials and synthesis of results

Only trials evaluated as high or unclear risk of bias were available for meta-analysis. Although the reasons for the high-risk components provided concern for potential bias, critical evaluation of results from meta-analyses enabled an overview of the current evidence and strength of effect to be presented, which permitted tentative conclusions to be proposed to advance research. Exploration of inter-trial compatibility of outcomes and assessment points identified back pain as the only possible comparison for exercise versus behavioural interventions, at 6 months, 12 months and 2 years. Reporting ‘mean change from baseline’ (SD)³³ for back pain intensity during the previous week on a 0–100 mm

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### Table 3  Summary assessment of the overall risk of bias for each trial

<table>
<thead>
<tr>
<th>Trial (authors, year, country)</th>
<th>Components of risk of bias/key risk criteria</th>
<th>Summary within trial</th>
<th>Comments on high-risk components</th>
</tr>
</thead>
<tbody>
<tr>
<td>Abbott et al (2010)³³</td>
<td>L L L L U U U</td>
<td>Unclear (3)</td>
<td>Low (4)</td>
</tr>
<tr>
<td>Christensen et al (2003)¹³</td>
<td>U U U U U U H</td>
<td>High (1)</td>
<td>Unclear (6)</td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td>One high-risk component: 6</td>
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<tr>
<td></td>
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<td></td>
<td>No primary outcome measure specified</td>
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<td>No primary end point specified</td>
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<td>No ITT reported</td>
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<td></td>
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<td></td>
<td>Poor reporting, lacking detail across all components</td>
</tr>
<tr>
<td>Sogaard et al (2006)³²</td>
<td>U L L L U U H</td>
<td>High (1)</td>
<td>Unclear (3)</td>
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<td></td>
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<td>Low (3)</td>
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<td>Differences at baseline re gender</td>
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<td>Low (2)</td>
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<td></td>
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</table>

Components of risk of bias/key risk criteria: 1, sequence generation; 2, allocation concealment; 3, blinding of participants, personnel and outcome assessors; 4, incomplete outcome data; 5a, short-term selective outcome reporting; 5b, long-term selective outcome reporting; 6, other potential threats to validity. Levels of risk of bias: H, high risk of bias; U, unclear risk of bias; L, low risk of bias. Summary WITHIN a study: L, low risk of bias for all key risk criteria; U, unclear risk of bias for one or more key risk criteria; H, high risk of bias for one or more key risk criteria.

ITT, intention-to-treat.
Visual Analogue Scale was evaluated as a comparable outcome to reporting median (range) for ‘mean back pain intensity’ within the previous 14 days on a 0–10 scale as part of the Low Back Pain Rating Scale.

At 6 months, the evidence from one trial suggested that intervention might reduce back pain, with a behavioural intervention being beneficial compared with an exercise intervention (figure 2). The pooled random effects (0.72, 95% CI –0.25 to 1.69) did not support evidence of an effect at 6-month short term.

At 12 months, the evidence from one trial suggested that intervention might reduce pain, with a behavioural intervention being beneficial compared with an exercise intervention (figure 3). The pooled random effects (0.52, 95% CI –0.45 to 1.49) did not support evidence of an effect at 12-month long term.

At 2 years, the evidence from one trial suggested that intervention might reduce pain, with a behavioural intervention being beneficial compared with an exercise intervention (figure 4). The pooled random effects (0.75, 95% CI –0.46 to 1.96) did not support evidence of an effect at 2-year long term. Overall, there was no evidence that intervention changes pain in the shorter (6 months) or long term (12 months or 2 years).

Additional analyses
The wide CIs for pooled effects indicated intervention could be potentially beneficial or harmful. No evidence from supportive analyses conflicted with the primary analyses.

DISCUSSION
Summary of evidence
Evidence was assessed from two randomised controlled trials (188 participants) conducted across two countries focused to lumbar fusion as a consequence of predominantly degenerative causes. Both trials included individualised 1:1 management and one trial included group management. Interventions were grouped into exercise versus behavioural comparisons. One trial was evaluated as high risk of bias and one as unclear. There were multiple issues contributing to the high risk of bias for Christensen et al. The number of issues did lessen in subsequent reporting of the Christensen trial, suggesting that trial reporting was problematic. One hundred and fifty-nine participants from the two trials were included in the meta-analyses. No patients >65 years were included in either of the two trials. This is problematic with a documented increase in patients undergoing lumbar spinal fusion from 2000/2001 to 2009/2010 in patients in the UK aged 60–74 years of 14%–22% and aged 75+ years of 2%–6%.

Results from the trial by Abbott et al, which was rated as unclear risk of bias, indicated no statistically significant difference across groups (n=54, n=53) with regard to pain, in the short or long term. Although findings from Christensen et al indicated a statistically significant benefit of behavioural intervention reducing pain in the short and longer term, there were multiple issues contributing to the high risk of bias for this trial and low numbers of participants (n=29, n=26 and n=26 in the three groups). It is noted that for Christensen et al, the 95% CIs were positioned completely within the ‘favour’ of behavioural intervention, illustrating some conflict within the pooled evidence from the two trials. Also of note were the narrower CIs for Abbott et al reflecting the much larger sample size in that trial. The pooled random effects of results from the two trials provided no supporting evidence of an effect. Overall, there was no evidence that physiotherapy management changes back pain.

The strengths of this review are its focus to physiotherapy intervention and it being the first in this area, exploring the breadth of potential physiotherapy interventions. It is, therefore, not possible to compare the
findings with previous reviews. The key findings of this review (only two trials, no evidence of the benefits of physiotherapy management) are a concern owing to the documented re-operation rate, re-hospitalisation rate and increasing numbers of patients undergoing lumbar spinal fusion surgery. When combined with existing literature reporting variability of conclusions and minimal data on the reported success for patient outcome following lumbar fusion, this concern contributes to a developing health problem, with available data supporting 25% patients reporting no change or worsened pain following surgery and 40% patients reporting dissatisfaction regarding their outcome 12 months following surgery. Consequently, there are major societal and economic implications from this ongoing disability, dissatisfaction and requirement for further intervention. Effective rehabilitation of patients following lumbar spinal fusion surgery is therefore an important issue.

Limitations
Differences were evident in the content and nature of interventions, selection of outcome measures and timing of assessment points, contributing to heterogeneity in treatment effects. Differences in components of the physiotherapy interventions might be explained by diversity in practice between countries. These differences limited the possible comparisons in the meta-analysis. Surprisingly, it was not possible to conduct an analysis of occupational outcome measures, despite their importance being identified within the literature.

Considerable heterogeneity was present in the evidence for behavioural intervention for pain at 6 months (I² 88%), 12 months (I² 88%) and 2 years (I² 92%), perhaps explaining no evidence of an effect for all evaluations. This anticipated heterogeneity was accounted for by using the random effects model.

Using GRADE (the Grading of Recommendations Assessment, Development and Evaluation system), the quality of the body of evidence for physiotherapy rehabilitation in the management of patients following lumbar fusion, based on the two trials included in the meta-analyses of behavioural versus exercise intervention, is very low for back pain in both the short and long term. These estimates are interpreted as ‘little confidence in the effect estimate: the true effect is likely to be substantially different from the estimate of effect’. Downgrading of quality was due to high risk of bias and issues of imprecision and inconsistency. The conclusions of cost-effectiveness in the subsequent reporting of Christensen et al are uncertain considering the very low quality.

The few trials available of high and unclear risk of bias and considerable heterogeneity illustrate the necessity for a high-quality and properly powered trial to evaluate a post-lumbar spinal fusion population. Meta-analyses depend upon the availability of published trials, and in this topical area that is increasing in profile, only two trials were available. The very low quality of existing trials is consistent with earlier findings for physiotherapy management post-lumbar discectomy. Within physiotherapy, there is currently limited scope for good quality meta-analyses with well-reported and rigorous criteria for trial inclusion. Owing to the limited comparability of outcome measures possible, consensus for a minimum core set of outcome measures for specific populations is also required.

CONCLUSIONS
This systematic review has identified inconclusive very low-quality evidence for the effectiveness of physiotherapy management following lumbar spinal fusion. Best practice for physiotherapy management is, therefore, unclear. There is no identified potential benefit for improving pain. However, the limited comparability of outcome measures and retrieval of only two trials reflects a lack of published research in this area. This gap requires urgent consideration, in a properly powered clinical trial with attention to quality and, in particular, to ensuring a low risk of bias. Inclusion of other outcomes that are important to the WHO’s classification is important, particularly occupational outcomes.

Contributors AR and GE are senior lecturers in physiotherapy and NH, RB and GJ are lecturers in physiotherapy. E-JP is a clinical physiotherapist. CW has long-standing professional interests in the quality and reporting of randomised controlled trials in physiotherapy. AR, GE, E-JP, NH and GJ have a professional focus to musculoskeletal physiotherapy. AR and CW were responsible for the conception of the study. All authors have contributed to the systematic review and have been involved in developing the content of the article. AR wrote the first draft of the paper and developed it initially with CW. AR has worked with all authors reworking content into subsequent drafts. All authors gave final approval of the version to be published. AR is the guarantor.

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Physiotherapy rehabilitation following lumbar spinal fusion


Physiotherapy rehabilitation following lumbar spinal fusion: a systematic review and meta-analysis of randomised controlled trials

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