

## Further support for the psychometric properties of the Farsi version of Perth Alexithymia Questionnaire

Lashkari, Arezou; Dehghani, Mohsen; Sadeghi-Firoozabadi, Vahid; Heidari, Mahmood; Khatibi, Ali

DOI:

[10.3389/fpsyg.2021.657660](https://doi.org/10.3389/fpsyg.2021.657660)

License:

Creative Commons: Attribution (CC BY)

*Document Version*

Publisher's PDF, also known as Version of record

*Citation for published version (Harvard):*

Lashkari, A, Dehghani, M, Sadeghi-Firoozabadi, V, Heidari, M & Khatibi, A 2021, 'Further support for the psychometric properties of the Farsi version of Perth Alexithymia Questionnaire', *Frontiers in Psychology*, vol. 12, 657660. <https://doi.org/10.3389/fpsyg.2021.657660>

[Link to publication on Research at Birmingham portal](#)

### General rights

Unless a licence is specified above, all rights (including copyright and moral rights) in this document are retained by the authors and/or the copyright holders. The express permission of the copyright holder must be obtained for any use of this material other than for purposes permitted by law.

- Users may freely distribute the URL that is used to identify this publication.
- Users may download and/or print one copy of the publication from the University of Birmingham research portal for the purpose of private study or non-commercial research.
- User may use extracts from the document in line with the concept of 'fair dealing' under the Copyright, Designs and Patents Act 1988 (?)
- Users may not further distribute the material nor use it for the purposes of commercial gain.

Where a licence is displayed above, please note the terms and conditions of the licence govern your use of this document.

When citing, please reference the published version.

### Take down policy

While the University of Birmingham exercises care and attention in making items available there are rare occasions when an item has been uploaded in error or has been deemed to be commercially or otherwise sensitive.

If you believe that this is the case for this document, please contact [UBIRA@lists.bham.ac.uk](mailto:UBIRA@lists.bham.ac.uk) providing details and we will remove access to the work immediately and investigate.



# Further Support for the Psychometric Properties of the Farsi Version of Perth Alexithymia Questionnaire

Arezou Lashkari<sup>1</sup>, Mohsen Dehghani<sup>1\*</sup>, Vahid Sadeghi-Firoozabadi<sup>1</sup>, Mahmood Heidari<sup>1</sup> and Ali Khatibi<sup>2,3\*</sup>

<sup>1</sup> Department of Psychology, Shahid Beheshti University, Tehran, Iran, <sup>2</sup> Centre of Precision Rehabilitation for Spinal Pain (CPR Spine), School of Sport, Exercise and Rehabilitation Sciences, College of Life and Environmental Sciences, University of Birmingham, Birmingham, United Kingdom, <sup>3</sup> Centre for Human Brain Health, University of Birmingham, Birmingham, United Kingdom

## OPEN ACCESS

### Edited by:

Paola Gremigni,  
University of Bologna, Italy

### Reviewed by:

Sina Hafizi,  
University of Manitoba, Canada  
Rory Allen,  
Goldsmiths University of London,  
United Kingdom

### \*Correspondence:

Mohsen Dehghani  
m.dehghani@sbu.ac.ir  
Ali Khatibi  
Ali.khatibi@gmail.com

### Specialty section:

This article was submitted to  
Quantitative Psychology  
and Measurement,  
a section of the journal  
Frontiers in Psychology

**Received:** 23 January 2021

**Accepted:** 23 March 2021

**Published:** 14 April 2021

### Citation:

Lashkari A, Dehghani M, Sadeghi-Firoozabadi V, Heidari M and Khatibi A (2021) Further Support for the Psychometric Properties of the Farsi Version of Perth Alexithymia Questionnaire.  
*Front. Psychol.* 12:657660.  
doi: 10.3389/fpsyg.2021.657660

Alexithymia is defined as the lack of words to describe emotions and is associated with different psychopathologies. Various tools have been developed for measuring alexithymia; each has its limitations. A new questionnaire, Perth Alexithymia Questionnaire (PAQ), was developed to simultaneously assess positive and negative dimensions. Validation of such a tool in different cultures allows cross-cultural health psychology studies and facilitates knowledge transfer in the field. We aimed to examine the psychometric features of the PAQ in the Farsi-speaking population in Iran. Four-hundred-twenty-nine university students were asked to complete the PAQ, the Toronto Alexithymia Scale (TAS-20), Beck Depression Inventory (BDI-II), Beck Anxiety Inventory (BAI), and emotion regulation questionnaire (ERQ). Concurrent validity, discriminant validity, internal consistency, and test-retest reliability and factor structure were investigated. Confirmatory factor analysis showed a five-factor model identical to the original questionnaire. The questionnaire indicated good internal consistency ( $0.82 < \alpha < 0.94$ ). Test-retest reliability was acceptable for all subscales. The correlations between PAQ and its subscales with BDI-II, BAI, and TAS, and expression suppression subscale of ERQ were strong for concurrent validity. Concerning the discriminant validity, PAQ and its subscales were not correlated with reappraisal subscales of ERQ. The present findings suggest that the Farsi version of PAQ has strong psychometric properties and is appropriate for use in the Farsi-speaking population.

**Keywords:** alexithymia, psychometric properties, validity, reliability, Farsi (Persian)

## INTRODUCTION

Sifneos (1973) introduced alexithymia by describing psychosomatic patients who could not find appropriate words to express their emotions. Psychosomatic disorders have been described in various forms in previous versions of the Diagnostic and Statistical Manual of Mental Disorders (DSM). DSM-I described psychosomatic disorders in a section called psychophysiological autonomic and visceral disorders. DSM 5 refers to psychosomatic disorders in Somatic Symptom and Related Disorders (Moldovan et al., 2015). In all disorders of somatic symptom and related disorders, physical symptoms along with distress and impairment are prominent (American Psychiatric Association [APA], 2013). Alexithymia is one of the salient features of psychosomatic disorders. These patients show somatic symptoms instead of expressing emotions. Although alexithymia was coined first as a standard feature in psychosomatic disorders, it is

prevalent in the normal population (Honkalampi et al., 2000; Kokkonen et al., 2001). Most researchers proposed that alexithymia is related to psychological pathologies. For example, Honkalampi et al. (2001) stated that the severity of depression was associated with alexithymia. de Bruin et al. (2019) showed that substance use disorder (SUD) patients with post-traumatic stress disorder (PTSD) are more alexithymic than SUD patients without PTSD. Iskríc et al. (2020) suggest that alexithymia is related to non-suicidal self-injury and suicidal ideation. Passardi et al. (2019) indicated that alexithymia played an essential role in facial emotion recognition of negative emotions in PTSD. In addition, McCallum et al. (2003) showed that alexithymia could affect the treatment outcomes by interfering with the improvement of general symptoms.

These studies demonstrated the importance of alexithymia as a construct. Toronto Alexithymia Scale (TAS) (Bagby et al., 1994) and Bermond-Vorst Alexithymia Questionnaire (BVAQ) (Vorst and Bermond, 2001) are the two most commonly used measures. TAS includes difficulty describing feelings (DDF), difficulty identifying feelings (DIF), and externally oriented thinking (EOT). TAS-20 is one of the first developed measures of alexithymia for both clinical and non-clinical samples. However, some researchers claimed that the EOT subscale of TAS does not have enough internal consistency (Kojima et al., 2001; Taylor et al., 2003; Bagby et al., 2020), and items load poorly on the related latent factor (Preece et al., 2018a,b). TAS correlated positively with negative affect measures and raised concern that TAS might assess distress rather than alexithymia construct (Preece et al., 2020). The other criticism of TAS is that it does not take the valence of emotions into account (Preece et al., 2018b).

Bermond-Vorst Alexithymia Questionnaire, another commonly used tool, assesses cognitive and affective dimensions of alexithymia. The cognitive dimension of alexithymia refers to the processing of emotions at a cognitive level. The affective dimension measures the subjective report of individual emotional experience. Emotionalizing and fantasizing characterize affective alexithymia. Alexithymia's cognitive dimension involve analyzing, identifying, and verbalizing emotions (Bermond et al., 2007). Differentiating the two dimensions of alexithymia is considered controversial (Goerlich, 2018). Some studies have failed to show the ability to differentiate between the two dimensions (Watters et al., 2016). Some researchers argued that difficulty in emotionalizing is not a valid indicator of alexithymia, as it does not distinguish between negative and positive reactivity to emotions (Preece et al., 2017).

Perth Alexithymia Questionnaire (PAQ; Preece et al., 2018b) was developed to address the earlier tool's limitations. PAQ is a 24-item questionnaire containing 5-subscales: positive DIF (P-DIF), negative DIF (N-DIF), positive DDF (P-DDF), negative DDF (N-DDF), and general EOT (G-EOT). PAQ is developed based on the attention-appraisal model of emotions (Gross, 2015). The model contains four stages: situation, attention, evaluation, and response. When an emotional response becomes the stimulus (situation stage) that is the target of valuation, the person pays attention to it (attention stage) and starts to evaluate it (appraisal). Then, the individual might respond to it (response stage). According to the model, alexithymic people cannot focus

their attention on emotional response. Furthermore, individuals with high levels of alexithymia cannot evaluate the emotional response as what it is and what it means. As a result, DIF and DDF are indicators of the appraisal stage. The salient point is that the appraisal stage can consist of valence. Both DIF and DDF subscales of PAQ comprise positive and negative valences and used to calculate the difficulty in the appraisal feeling (DAF) composite subscale. By combining the 5-subscale, composite subscales can be produced. In this regard, positive difficulty appraising feeling (P-DAF) includes positive DIF and positive DDF. Negative difficulty appraising feeling (N-DAF) is the sum of negative DIF and negative DDF (Preece et al., 2018b).

As mentioned, PAQ is constructed based on the attention-appraisal model of emotions. Most studies revealed that alexithymia reflects a deficit in regulating emotions (Swart et al., 2009; Pandey et al., 2011). Also, it has been suggested that alexithymic individuals used more suppression and less reappraisal for emotion regulation (Swart et al., 2009). Cognitive reappraisal means cognitive changes in the affective influence of the emotion-eliciting situation. Expressive suppression inhibits emotional behavior. Cognitive reappraisal is assumed as an adaptive emotion regulation strategy. However, expressive suppression is indicative of a maladaptive emotion regulation strategy (Gross and John, 2003). It is shown that emotional expression is strongly related to the cultural context (Altarriba and Kazanas, 2017). So, it is crucial to assess the psychometric features of the measures in other cultures.

Research showed the role of culture in factorial structure (Fernández-Jiménez et al., 2013). TAS-20 and more recently PAQ are measures for assessing the alexithymia in Iran. As we know, EOT component of the Farsi version of TAS-20 has lower internal consistency (Besharat, 2008). Besides, other studies aimed at validating PAQ in Iran failed to recruit a big enough sample or to report confirmatory analysis (Heydari et al., 2020; Mousavi Asl et al., 2020). Accordingly, the current study aimed to examine the validity and reliability of PAQ in Iran with specific reference to its internal consistency, test-retest reliability, concurrent and discriminant validity, and factor structure. In this regard, depression, anxiety inventories and TAS, suppression subscale of emotion regulation questionnaire (ERQ) were used for assessing concurrent validity. Discriminant validity was assessed by the cognitive reappraisal subscale of the ERQ. Confirmatory factor analyses of PAQ tested the construct validity. Considering the psychometric properties of the original version of the PAQ, we expected the measure to be highly reliable. We hypothesized that PAQ is an excellent measure to assess alexithymia as reflected in its correlation with other measures of emotional and mood difficulties. We also expected the PAQ to have a unique contribution above that of the previous measures as reflected in factor analysis.

## Method

### Participants

Participants were recruited via advertisements on notice boards in the universities. They were included if they reported a good level of literacy, including reading and understanding Farsi.

Those with a history of traumatic experiences in the previous 6 months to the data collection point were excluded. A total of 436 individuals answered the questionnaires. Incomplete questionnaires were removed from the final sample leading to 429 (63.2% female, mean age =  $21.35 \pm 2.88$ , range = 18–39) unique responses that were included in the statistical analyses. All participants were university students, including 364 undergraduate, 41 masters, 11 Ph.D. students. The study was approved by the Research Ethics Committee of Shahid Beheshti University, Tehran, Iran.

## MATERIALS AND METHODS

### Perth Alexithymia Questionnaire

Perth Alexithymia Questionnaire (Preece et al., 2018b) consists of 24 self-report measures rated on a seven-point Likert scale (1 = strongly disagree; 7 = strongly agree), with high scores demonstrating high levels of alexithymia. PAQ, the only alexithymia measure including the valence of emotions, consists of five subscales: negative-difficulty identifying feelings (N-DIF), positive-difficulty identifying feelings (P-DIF), negative-difficulty describing feelings (N-DDF), positive-difficulty describing feelings (P-DDF), and general-externally orientated thinking (G-EOT). “When I am feeling bad, I can’t tell whether I’m sad, angry or scared” and “when I’m feeling good, I get confused about what emotion it is” are the samples of N-DIF and P-DIF. The subscales can combine and produce the composite subscales. PAQ demonstrated good concurrent and discriminant validity and good internal consistency (Cronbach’s alpha of the subscales ranged from 0.87 to 0.96) (Preece et al., 2018b).

We followed a standard translation and back-translation procedure for the introduction of the Farsi version of the PAQ. First, the original questionnaire was translated into Farsi, and then a bilingual professional psychologist back-translated it into English. The back-translated version was compared with the original English one, a few minor corrections were applied and the final version of the PAQ, used in this study, was reached.

### 20-Item Toronto Alexithymia Scale

20-Item Toronto Alexithymia Scale (Bagby et al., 1994) is a 20-item measure consisting of three subscales: difficult identifying feelings (DIF), DDF, and EOT style. Each item is rated on a 5-point Likert scale (1 = strongly disagree to 5 = strongly agree). “When I am upset, I don’t know if I am sad, frightened or angry”, “It is difficult for me to find the right words for my feelings”, and “Being in touch with emotions is essential” are the samples of DIF, DDF, and EOT, respectively. Five key items are reverse scored (4, 5, 10, 11, and 19). TAS-20 demonstrated good internal consistency (Cronbach’s alpha ranged from 0.66 to 0.75) and test-retest reliability (0.77). The three-factor structure was congruent with the theoretical model underlying TAS-20 (Bagby et al., 1994). TAS psychometric properties have been tested and confirmed by various studies (Bressi et al., 1996; Taylor et al., 2003; Säkkinen et al., 2007). The Farsi version of TAS demonstrated good validity and reliability in the Farsi-speaking population (Besharat, 2008).

### Emotion Regulation Questionnaire

Emotion regulation questionnaire (Gross and John, 2003) is a short questionnaire designed to separate two subscales: expressive suppression (I keep my emotions to myself) and cognitive reappraisal (When I want to feel more positive emotion [such as joy or amusement], I change what I’m thinking about). ERQ is answered on a 7-point Likert scale (1 = strongly disagree, 7 = strongly agree). The Cronbach’s alpha was found to be in the range of 0.68–0.82 in different populations. ERQ demonstrated good convergent and discriminant validity (Gross and John, 2003). ERQ showed good validity and reliability in the Farsi-speaking population. The Cronbach’s alpha coefficient for cognitive reappraisal and expressive suppression were 0.78 and 0.60, respectively (Ghasempour et al., 2012).

### Beck Depression Inventory-II

Beck depression inventory-II (Beck et al., 1996) is a widely used measure for assessing depression in both clinical and non-clinical populations. BDI-II contains 21 items, each answer being scored on a Likert scale value of 0–3 in which higher scores indicate the existence of depression symptoms. One of the scale items is as follows: I do not feel sad, I feel sad much of the time, I am sad all the time, I am so sad or unhappy that I can’t stand it. Individuals are asked to select one of four possible items in each question based on the last 2 weeks state. The internal consistency is reported around 0.90, and test-retest reliability from 0.73 to 0.96. BDI-II demonstrated two-factor of cognitive-affective and somatic-vegetative (Wang and Gorenstein, 2013). In the Farsi-speaking sample, the same two-factor model is confirmed (Ghassemzadeh et al., 2005), and internal consistency and test-retest reliability were 0.87 and 0.74.

### Beck Anxiety Inventory

The BAI (Beck et al., 1988) is a 21-item self-report questionnaire that asks about common anxiety symptoms such as being scared, nervous, and unsteady. The items are rated on a 4-point Likert scale (0 = not at all, 3 = severely). The total score ranged from 0 to 63. The higher scores indicate a higher level of anxiety. BAI has the clinical classification, with scores from 0 to 7 displaying minimal anxiety, 8–15 as mild anxiety, 16–25 as moderate anxiety, and 26–63 as severe anxiety. Scores greater than 16 indicate clinically significant anxiety (Beck et al., 1988). BAI internal consistency (alpha Cronbach) and test-retest reliability reported as 0.91 and 0.65, respectively. BAI demonstrated a two-factor solution and concurrent validity with other anxiety measures and the BDI-II as well (Bardhoshi et al., 2016). In the Iranian population, the internal consistency coefficient (alpha Cronbach) was around 0.90, and the three-factor model is confirmed (Dobson and Mohamadkhani, 2008).

### Data Analysis

Confirmatory factor analyses (CFA) was performed using LISREL 8.80. All other analyses, such as correlations and descriptive statistics, were done by SPSS 24. As mentioned

before, the PAQ concept is based on the attention-appraisal model of alexithymia. As a result, the subscale of PAQ can combine and produce composite subscales according to the theory (Preece et al., 2018b). Five theoretical models for the factor structure have been suggested. In this study, CFA was applied for all five models. CFA was conducted using the maximum likelihood estimation based on the Pearson covariance matrix. Maximum likelihood requires that the data display both univariate and multivariate normality, despite the robustness to violation of normality as long as the sample size is large (Gerbing and Anderson, 1985; Finney and DiStefano, 2006).

The fitness of the model was evaluated with the most important indices such as root mean square error of approximation (RMSEA), comparative fit index (CFI), normed fit index (NFI), incremental fit index (IFI), standard root mean squared residual (SRMR), Akaike information criterion (AIC), and relative chi-square ( $\chi^2/df$ ). Acceptable fit values of CFI, NFI, and IFI are greater than 0.90, while the acceptance value of  $\chi^2/df$  is less than 5 (Hu and Bentler, 1999; Hooper et al., 2008). The acceptable value of RMSEA and SRMR is less than 0.08 and 0.1, respectively (Schumacker and Lomax, 2016). AIC is a criterion to compare different models, and lower values demonstrate better fitness (Kline, 2015).

Cronbach's alpha was calculated to assess internal consistency. The values greater than 0.90 indicate excellent consistency, while those greater than 0.80 and 0.70 indicate good and acceptable internal consistency, respectively (Groth-Marnat, 2009). The test-retest reliability was examined in a separate group of 59 participants (76.3% female) with a mean age of 27.49 (SD = 5.12, range = 18–47) that were recruited only for this purpose. They completed the PAQ questionnaire two times with a 2-week interval between them. The test-retest reliability was quantified using a Pearson correlation estimate for total scores.

The Pearson correlation coefficient was used to assess concurrent and discriminant validity. We hypothesized that a medium to large (around 0.3 and 0.5) correlation (Cohen, 1998) between PAQ and TAS, BDI-II, BAI, and expressive suppression subscale of ERQ indicates concurrent validity. We expected a small, around 0.1 (Cohen, 1998) and even negative correlation of PAQ and cognitive reappraisal subscale of ERQ as discriminant validity.

## Procedure

At the beginning of the session, participants received the consent form. After reading and accepting to participate, they received the battery of questionnaires and were instructed to complete them by paying attention to the instruction on top of each questionnaire. The researcher (AL) was accessible during the time they were completing the questionnaires. At the end of the session, participants were debriefed. For the test-retest, due to the restriction imposed as the consequence of the COVID19 pandemic, a separate group of participants received a link with the online format of the translated PAQ questionnaire. They received a similar link 2 weeks later with the same questionnaire and were asked to answer it within 24 h.

## RESULTS

### Descriptive Statistics

**Table 1** presents the mean and standard deviation for subscales and total PAQ, and other measures in the study. The minimum and maximum mean and standard deviation belonged to the P-DIF and the G-EOT subscales, respectively ( $M = 9.84$ ,  $SD = 4.96$ ;  $M = 21.38$ , and  $SD = 9.57$ ).

### Confirmatory Factor Analyses

Before applying the CFA, all 24 items were checked for normal distribution. A skewness value of  $\pm 1$  and  $\pm 2$  is considered excellent and acceptable, respectively. While a value of  $\pm 3$  is described as highly skewed. Kurtosis greater than 10 indicates non-normal distribution (George and Mallery, 2006; Kline, 2015). In the present study, the skewness ranged between 0.24 and 1.52, and the range of kurtosis was from  $-1.15$  to 1.47, indicating a normal distribution of the scores.

Model 1 was a one-factor model in which all 24 items were loaded in general alexithymia. Model 2 was a two-factor model in which items were loaded on G-EOT and G-DAF. This model discriminates attention and appraisal stages of emotion evaluation. Model 3 was a three-factor model consisting of

**TABLE 1** | Descriptive statistics for the administered measures.

Measure/subscale		Total	
		M	SD
Perth Alexithymia Questionnaire	P-DIF	9.84	4.96
	N-DIF	12.40	5.91
	P-DDF	11.05	5.83
	N-DDF	13.32	6.49
	G-EOT	21.38	9.57
	G-DIF	22.25	9.96
	G-DDF	24.37	11.27
	N-DAF	25.72	11.72
	P-DAF	20.89	10.27
	G-DAF	46.62	20.28
TAS	ALEXI	68.01	27.05
	DIF	15.71	5.69
	DDF	13.27	4.34
ERQ	EOT	18.94	3.90
	Expressive and suppression	13.23	5.42
	Cognitive and reappraisal	26.34	7.33
BAI		11.86	9.49
BDI-II		14.62	10.488

*M*, mean; *SD*, standard deviation; *PAQ*, Perth Alexithymia Questionnaire; *N-DIF*, negative-difficulty identifying feelings; *P-DIF*, positive-difficulty identifying feelings; *NDDF*, negative-difficulty describing feelings; *P-DDF*, positive-difficulty describing feelings; *G-EOT*, general-externally orientated thinking; *GDIF*, general-difficulty identifying feelings; *G-DDF*, general-difficulty describing feelings; *N-DAF*, negative-difficulty appraising feelings; *PDAF*, positive-difficulty appraising feelings; *DAF*, general-difficulty appraising feelings; *ALEXI*, alexithymia; *BAI*, beck anxiety inventory; *BDI*, beck depression inventory; *TAS*, Toronto alexithymia questionnaire; *DIF*, difficulty identifying the feeling; *DDF*, difficulty describing the feeling; *EOT*, externally oriented thinking; and *ERQ*, emotion regulation questionnaire.

G-DIF, G-DDF, and G-EOT. G-DIF is obtained by combining positive and negative DIF, and the G-DDF is obtained by summing the positive DDF and negative DDF. There was no distinction between the valence of components in this model. Model 4 is a three-factor model based on the distinction of valence and includes G-EOT, N-DAF, and P-DAF factors. N-DAF is referred to as negative DIF, and negative DDF and P-DAF are created by collecting positive DIF and positive DDF. Model 5 is a five-factor model, and all factors were divided based on valence, DIF, and DDF. The items were loaded on G-EOT, N-DIF, P-DIF, N-DDF, and P-DDF. The goodness of fit values in **Table 2** indicates that models 5 and 1 revealed the best and poorest fit indices, respectively. All the values of model 5 demonstrate that model 5 is the best-fitted model comparing other models. All items were loaded well on five hypothesized factors (**Figure 1**).

## Reliability

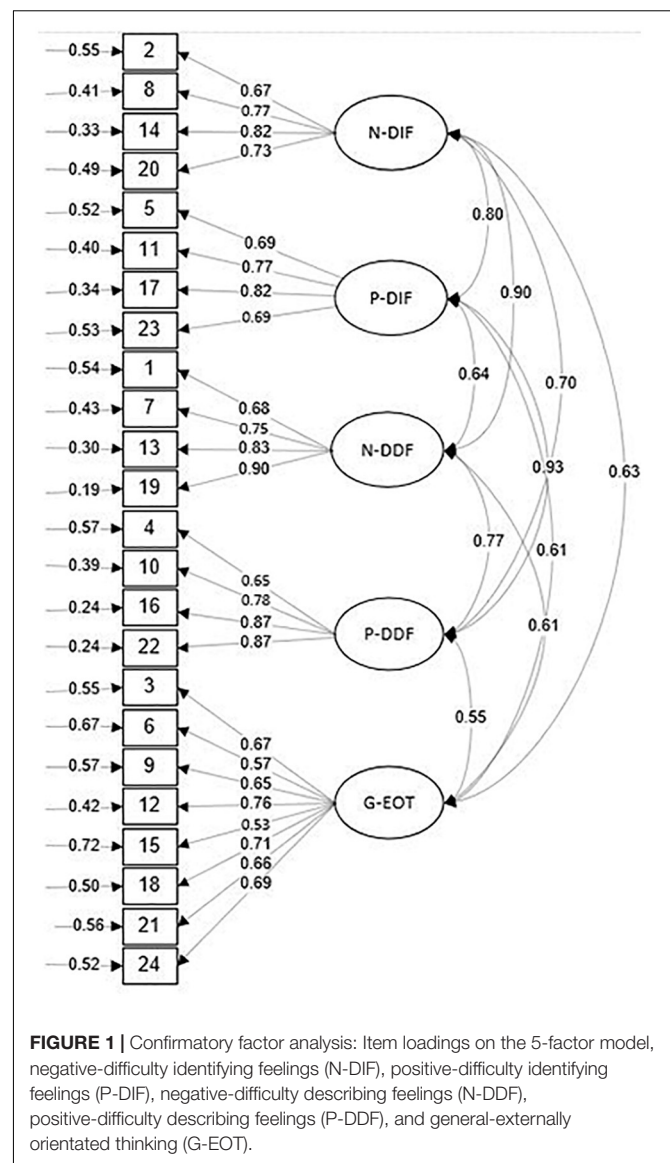
We examined the reliability of the measure in two different ways. First, we measured the test-retest reliability for the subscales and total scales, which are presented in **Table 3** (left column). Test-retest reliability of all subscales was statistically significant after at least 2 weeks ( $r > 0.69$ ). Cronbach's alpha coefficient ranged from 0.82 to 0.94 for subscales and 0.94 for the total score, suggesting the good internal consistency of PAQ.

## Concurrent and Discriminant Validity

**Table 4** presents the correlation of all subscales of PAQ and TAS, ERQ, BAI, and BDI-II. As expected, the correlations between all subscales of PAQ and TAS are statistically significant, ranging from 0.18 to 0.73. Also, there are positive correlations between PAQ and BDI-II and BAI. The expressive suppression subscale of ERQ and PAQ subscales were significantly correlated (0.37–0.53), and no significant correlation was found between PAQ and cognitive reappraisal.

## DISCUSSION

This study aimed to investigate the utility of PAQ in Farsi-speaking people living in Iran. We aimed to determine which of the original model's five variants would fit best in the Farsi-speaking population. Although most of the fit indices were



acceptable in all five models, model 5 (which consists of five-factors) was the best. In model 5, all 24 items were loaded on the related factors. Therefore, it can be claimed that the Farsi version of the PAQ is a multidimensional questionnaire. This claim is consistent with the results of Preece et al. (2018b). As assumed by prior studies (Barrett et al., 2001), differences between positive and negative emotions are important for emotion regulation because the individuals differ in their emotional experiences. One of the most distinctive features of the PAQ is that it considers the valence of emotions, which is well described in separating the positive and negative valence in appraising factors. P-DIF, N-DIF, P-DDF, and N-DDF are the indicators of the appraisal stage of the model. Model 5 contains all four mentioned factors plus the EOT, which is not valence specified.

Testing the measure's reliability, we found the PAQ subscale and the composite subscales to be reliable constructs similar

**TABLE 2** | Goodness-of-fit index values from confirmatory factor analysis of the 24 Perth Alexithymia Questionnaire items.

Model	Chi-Square	Df	$\chi^2/df$	RMSEA	CFI	NFI	IFI	AIC	SRMR
Model1	2,857.29	252	11.33	0.15	0.91	0.90	0.91	2,953.29	0.09
Model2	1,892.36	251	7.53	0.12	0.94	0.93	0.94	1,990.36	0.06
Model3	1,928.67	249	7.74	0.12	0.94	0.93	0.94	2,030.15	0.06
Model4	10.56	249	4.24	0.08	0.97	0.95	0.97	1,158.67	0.04
Model5	784.53	242	3.24	0.07	0.97	0.96	0.97	900.53	0.04

*Df*, the degree of freedom;  $\chi^2/df$ , relative chi-square; *RMSEA*, root mean square error; *CFI*, the comparative fit index; *NFI*, normed fit index; *IFI*, incremental fit index; *SRMR*, standard root mean squared residual; and *AIC*, Akaike information criterion.

to the original version (Preece et al., 2018b). The purpose of creating the new measure was to overcome the limitations of previous tools, including the low reliability of the EOT subscale.

**TABLE 3 |** Test-retest reliability and Cronbach's alpha reliability coefficients for administered measures.

Measure/subscales	Test-retest correlation	Cronbach's alpha ( $\alpha$ )
PAQ		
P-DIF	0.71**	0.82
N-DIF	0.72**	0.83
P-DDF	0.69**	0.87
N-DDF	0.71**	0.87
G-EOT	0.74**	0.85
G-DIF	0.76**	0.88
G-DDF	0.77**	0.90
N-DAF	0.75**	0.91
P-DAF	0.76**	0.91
G-DAF	0.81**	0.94
ALEXI	0.85**	0.94

PAQ, Perth Alexithymia Questionnaire; N-DIF, negative-difficulty identifying feelings; P-DIF, positive-difficulty identifying feelings; NDDF, negative-difficulty describing feelings; P-DDF, positive-difficulty describing feelings; G-EOT, general-externally orientated thinking; GDIF, general-difficulty identifying feelings; G-DDF, general-difficulty describing feelings; N-DAF, negative-difficulty appraising feelings; PDAF, positive-difficulty appraising feelings; DAF, general-difficulty appraising feelings; ALEXI, alexithymia.

\*\* $p < 0.01$ .

**TABLE 4 |** Pearson correlations between the Perth Alexithymia Questionnaire, beck depression inventory, beck anxiety inventory, emotion regulation questionnaire, and Toronto Alexithymia Scale.

Measures/ subscales	BDI-II BAI		ERQ		TAS-		
			Cognitive reappraisal	Expressive suppression	DIF	DDF	EOT
PAQ Subscales							
N-DIF	0.39**	0.33**	0.03	0.39**	0.73**	0.57**	0.18**
P-DIF	0.28**	0.21**	-0.04	0.37**	0.62**	0.50**	0.28**
N-DDF	0.35**	0.20**	0.01	0.44**	0.58**	0.67**	0.19**
P-DDF	0.26**	0.16**	-0.07	0.41**	0.52**	0.58**	0.22**
G-EOT	0.34**	0.21**	0.02	0.53**	0.47**	0.49**	0.28**
Composites							
G-DIF	0.37**	0.30**	-0.01	0.41**	0.74**	0.59**	0.25**
G-DDF	0.34**	0.20**	0.00	0.46**	0.61**	0.69**	0.22*
N-DAF	0.39**	0.28**	0.02	0.44**	0.69**	0.66**	0.20**
P-DAF	0.28**	0.19**	-0.02	0.41**	0.60**	0.58*	0.26**
G-DAF	0.37**	0.26**	-0.04	0.46**	0.70**	0.67**	0.25**
ALEXI	0.40**	0.27**	-0.01	0.53**	0.70**	0.68**	0.28**

PAQ, Perth Alexithymia Questionnaire; N-DIF, negative-difficulty identifying feelings; P-DIF, positive-difficulty identifying feelings; NDDF, negative-difficulty describing feelings; P-DDF, positive-difficulty describing feelings; G-EOT, general-externally orientated thinking; GDIF, general-difficulty identifying feelings; G-DDF, general-difficulty describing feelings; N-DAF, negative-difficulty appraising feelings; PDAF, positive-difficulty appraising feelings; DAF, general-difficulty appraising feelings; ALEXI, alexithymia; BAI, beck anxiety inventory; BDI, beck depression inventory; TAS, Toronto alexithymia questionnaire; DIF, difficulty identifying the feeling; DDF, difficulty describing the feeling; EOT, externally oriented thinking; and ERQ, emotion regulation questionnaire.

\*\* $(p < 0.01)$ ; \* $(p < 0.05)$ .

Most studies presented that EOT component has low reliability in different cultures (Bressi et al., 1996; Kojima et al., 2001; Taylor et al., 2003). In the present study, all the alpha coefficients were greater than 0.80, reflecting good internal consistency, particularly EOT subscale, which is consistent with the study by Preece et al. (2018b). The original study did not assess the test-retest reliability (Preece et al., 2018b). All subscales, especially the total alexithymia, indicated good test-retest reliability, which means that the scale results are consistent at different time points. The test-retest reliability results in the present study are comparable with those reported by Mousavi Asl et al. (2020).

To test the concurrent and discriminant validity of the PAQ, we measured the correlations between the subscales and other previously validated measures, including the BDI-II, BAI, TAS, and ERQ. In agreement with the original study (Preece et al., 2018b), PAQ was strongly correlated with all TAS subscales, which is another alexithymia measure. A score for general difficulties in identifying and describing feelings will be obtained by combining the negative and positive subscales. These two subscales, along with the externally orientated thinking subscale, are the same subscales that TAS measures too. Therefore, a high correlation between the two measures was expected. The correlation of PAQ with BDI-II and BAI is in line with the original study (Preece et al., 2018b). Individuals with depression and anxiety seemed to have difficulty recognizing and describing their emotions. Honkalampi et al. (2001) stated that alexithymia is associated with depression. Moreover, Honkalampi et al. (2000) reported that the prevalence of alexithymia in individuals with mild depression (BDI > 9) was 32%. Another study showed that difficulties in describing and identifying feelings, changes with mood and recovery from depression and were associated with a decrease in alexithymia (Saarijärvi et al., 2001). Alexithymia is also associated with anxiety. Difficulties describing and identifying feelings are related to anxiety disorders such as generalized anxiety disorder (GAD), and the presence of alexithymia is related to higher levels of anxiety (Berardis et al., 2008).

The expressive suppression subscale of the ERQ was correlated positively with the PAQ subscales. However, PAQ was not correlated with the cognitive reappraisal subscale of ERQ. Cognitive reappraisal and expressive suppression are considered adaptive and maladaptive emotion regulation, respectively (Gross and John, 2003). As mentioned before, alexithymia is defined as a deficiency in expressing emotions. Accordingly, it is expected that alexithymia would be correlated with expressive suppression. In line with the lack of relationship between alexithymia and cognitive reappraisal, it can be stated that alexithymia is an obstacle to regulating emotions (Gross, 1998). It can be assumed that the alexithymia has concurrent and discriminant validity.

The present study, despite its advantages, also suffers from some limitations. Firstly, the study was carried out among university students. Therefore, generalizing the finding to other populations, especially clinical samples, should be done with caution. Secondly, the cutoff point was not assessed, and this will limit the use of the measure for clinical research. Hence, for future studies, it is recommended to evaluate the psychometric feature of the PAQ in other populations, especially clinical samples.

Understanding emotion and its contribution to wellbeing is important in health psychology research. Because of the universality of emotions, they can help us in cross-cultural studies. Adaptation of the existing tools to new languages and in different cultures can facilitate cross-cultural health psychology research (Rudell and Diefenbach, 2008). PAQ (Preece et al., 2018b) is one of the most widely used tools to test alexithymia. Previous efforts in adapting and validating this tool among the Farsi-speaking population suffer from low sample size and lack of model estimation. In the current study, we examined the psychometric properties of a new adaptation of the PAQ (please see **Supplementary Material**) in a large Farsi-speaking sample in Iran. In summary, the study findings show that the Farsi version of PAQ is a valid and reliable measure. PAQ seems to be a promising measure for identifying the deficiency of emotions in the Farsi-speaking sample.

## DATA AVAILABILITY STATEMENT

The raw data supporting the conclusions of this article will be made available by the authors, without undue reservation.

## REFERENCES

- Altarriba, J., and Kazanas, S. A. (2017). "Emotions and Expressions across Cultures," in *The International Encyclopedia of Intercultural Communication*, eds Y. Y. Kim and K. McKay-Semmler (Hoboken, NJ: John Wiley & Sons, Inc), 1–10. doi: 10.1002/9781118783665.ieicc0247
- American Psychiatric Association [APA]. (2013). *Diagnostic and statistical manual of mental disorders (DSM-5®)*. Washington, DC: American Psychiatric Association.
- Bagby, R. M., Parker, J. D. A., and Taylor, G. J. (2020). Twenty-five years with the 20-item Toronto Alexithymia Scale. *J Psychosom Res* 131, 109940. doi: 10.1016/j.jpsychores.2020.109940
- Bagby, R. M., Parker, J. D., and Taylor, G. J. (1994). The twenty-item Toronto Alexithymia Scale—I. Item selection and cross-validation of the factor structure. *J Psychosom Res* 38, 23–32. doi: 10.1016/0022-3999(94)90005-1
- Bardhoshi, G., Duncan, K., and Erford, B. T. (2016). Psychometric meta-analysis of the English version of the Beck Anxiety Inventory. *J Couns Dev* 94, 356–373. doi: 10.1002/jcad.12090
- Barrett, L. F., Gross, J., Christensen, T. C., and Benvenuto, M. (2001). Knowing what you're feeling and knowing what to do about it: Mapping the relation between emotion differentiation and emotion regulation. *Cogn Emot* 15, 713–724. doi: 10.1080/02699930143000239
- Beck, A. T., Epstein, N., Brown, G., and Steer, R. A. (1988). An inventory for measuring clinical anxiety: psychometric properties. *J Consult Clin Psychol* 56, 893. doi: 10.1037/0022-006x.56.6.893
- Beck, A. T., Steer, R. A., and Brown, G. K. (1996). *Manual for the beck depression inventory-II*. San Antonio, TX: Psychological Corporation.
- Berardis, D., Campanella, D., Nicola, S., Gianna, S., Alessandro, C., Chiara, C., et al. (2008). The Impact of Alexithymia on Anxiety Disorders: a Review of the Literature. *Curr Psychiatry Rev* 4, 80–86. doi: 10.2174/157340008784529287
- Bermond, B., Clayton, K., Liberova, A., Luminet, O., Maruszewski, T., Ricci Bitti, P. E., et al. (2007). A cognitive and an affective dimension of alexithymia in six languages and seven populations. *Cogn Emot* 21, 1125–1136. doi: 10.1080/02699930601056989
- Besharat, M. A. (2008). Psychometric characteristics of Persian version of the Toronto alexithymia scale-20 in clinical and non-clinical samples. *Iran J Med Sci* 33, 1–6.
- Bressi, C., Taylor, G., Parker, J., Bressi, S., Brambilla, V., Aguglia, E., et al. (1996). Cross validation of the factor structure of the 20-item Toronto Alexithymia Scale: an Italian multicenter study. *J Psychosom Res* 41, 551–559. doi: 10.1016/s0022-3999(96)00228-0
- Cohen, J. (1998). *Statistical power analysis for the behavioral sciences*. New Jersey, NJ: Lawrence Erlbaum association.
- de Bruin, P. M. J., de Haan, H. A., and Kok, T. (2019). The prediction of alexithymia as a state or trait characteristic in patients with substance use disorders and PTSD. *Psychiatry Res* 282, 1–6. doi: 10.1016/j.psychres.2019.112634
- Dobson, K., and Mohamadkhani, P. (2008). Psychometric properties of Beck Depression Inventory- II in a sample of people with depressive disorder. *IRJ* 39, 80–89.
- Fernández-Jiménez, E., Pérez-San-Gregorio, M. Á, Taylor, G. J., Michael Bagby, R., Ayeart, L. E., and Izquierdo, G. (2013). Psychometric properties of a revised Spanish 20-item Toronto Alexithymia Scale adaptation in multiple sclerosis patients. *Int J Clin Health Psychol* 13, 226–234. doi: 10.1016/s1697-2600(13)70027-9
- Finney, S. J., and DiStefano, C. (2006). Non-normal and categorical data in structural equation modeling. In G. R. Hancock & R. O. Mueller (Hrsg.). *Structural equation modeling: a second course*. Greenwich, CT: Information Age Publishing, 269–314.
- George, D., and Mallery, P. (2006). *IBM SPSS statistics 26 step by step: a simple guide and reference*. New York, NY: Taylor & Francis.
- Gerbing, D. W., and Anderson, J. C. (1985). The Effects of Sampling Error and Model Characteristics on Parameter Estimation for Maximum Likelihood Confirmatory Factor Analysis. *Multivariate Behav Res* 20, 255–271. doi: 10.1207/s15327906mbr2003\_2
- Ghasempour, A., Eilbeigi, R., and Hassanzadeh, S. (2012). "Psychometric properties of the emotional regulation questionnaire of Gross and John in an Iranian sample," in *The Sixth Congress on Mental Health of University Students*. Gilan-Iran.
- Ghassemzadeh, H., Mojtabai, R., Karamghadiri, N., and Ebrahimkhani, N. (2005). Psychometric properties of a Persian-language version of the Beck Depression Inventory—Second edition: BDI-II-PERSIAN. *Depress Anxiety* 21, 185–192. doi: 10.1002/da.20070
- Goerlich, K. S. (2018). The Multifaceted Nature of Alexithymia - A Neuroscientific Perspective. *Front. Psychol* 9:1614. doi: 10.3389/fpsyg.2018.01614
- Gross, J. J. (1998). The emerging field of emotion regulation: An integrative review. *Rev Gen Psychol* 2, 271–299. doi: 10.1037/1089-2680.2.3.271
- Gross, J. J. (2015). Emotion Regulation: Current Status and Future Prospects. *Psychol Inq* 26, 1–26. doi: 10.1080/1047840x.2014.940781

## ETHICS STATEMENT

The studies involving human participants were reviewed and approved by Department of Psychology, Shahid Beheshti University. The patients/participants provided their written informed consent to participate in this study.

## AUTHOR CONTRIBUTIONS

AL was involved in design, data collection, analysis, and writing. MD and AK was involved in design and writing. VS-F and MH was involved in design. All authors contributed to the article and approved the submitted version.

## SUPPLEMENTARY MATERIAL

The Supplementary Material for this article can be found online at: <https://www.frontiersin.org/articles/10.3389/fpsyg.2021.657660/full#supplementary-material>



- Gross, J. J., and John, O. P. (2003). Individual differences in two emotion regulation processes: implications for affect, relationships, and well-being. *J Pers Soc Psychol* 85, 348–362. doi: 10.1037/0022-3514.85.2.348
- Groth-Marnat, G. (2009). *Handbook of psychological assessment*. New Jersey, NJ: John Wiley & Sons.
- Heydari, S., Lajmiri, S., Azadyekta, M., Barzegar, M., and Arshadi, M. (2020). Reliability and validity of Perth alexithymia questionnaire (PAQ) and its relation with cognitive emotion regulation. *Iranian Journal of Rooyesh\_e\_Ravanshenasi* 9, 73–80.
- Honkalampi, K., Hintikka, J., Laukkanen, E., Lehtonen, J., and Viinamäki, H. (2001). Alexithymia and depression: a prospective study of patients with major depressive disorder. *Psychosomatics* 42, 229–234. doi: 10.1176/appi.psy.42.3.229
- Honkalampi, K., Hintikka, J., Tanskanen, A., Lehtonen, J., and Viinamäki, H. (2000). Depression is strongly associated with alexithymia in the general population. *J Psychosom Res* 48, 99–104. doi: 10.1016/s0022-3999(99)00083-5
- Hooper, D., Coughlan, J., and Mullen, M. (2008). Structural equation modelling: guidelines for determining model fit. *Electron J Bus Res Methods Vol 6*, 53–60.
- Hu, L. T., and Bentler, P. M. (1999). Cutoff criteria for fit indexes in covariance structure analysis: Conventional criteria versus new alternatives. *Struct Equ Model* 6, 1–55. doi: 10.1080/10705519909540118
- Iskric, A., Ceniti, A. K., Bergmans, Y., McInerney, S., and Rizvi, S. J. (2020). Alexithymia and self-harm: A review of nonsuicidal self-injury, suicidal ideation, and suicide attempts. *Psychiatry Res* 288, 1–15. doi: 10.1016/j.psychres.2020.112920
- Kline, R. B. (2015). *Principles and practice of structural equation modeling*. New York, NY: Guilford publications.
- Kojima, M., Frasure-Smith, N., and Lespérance, F. (2001). Alexithymia following myocardial infarction: psychometric properties and correlates of the Toronto Alexithymia Scale. *J Psychosom Res* 51, 487–495.
- Kokkonen, P., Karvonen, J. T., Veijola, J., Läksy, K., and Jokelainen, J. (2001). Perceived and sociodemographic correlates of alexithymia in a population sample of young adults. *Compr Psychiatry* 42, 471–476. doi: 10.1053/comp.2001.27892
- McCallum, M., Piper, W. E., Ogradniczuk, J. S., and Joyce, A. S. (2003). Relationships among psychological mindedness, alexithymia and outcome in four forms of short-term psychotherapy. *Psychol Psychother Theory Res Pract* 76, 133–144. doi: 10.1348/147608303765951177
- Moldovan, R., Radu, M., Baban, A., and Dumitrascu, D. L. (2015). Evolution of Psychosomatic Diagnosis in DSM. Historical Perspectives and New Development for Internists. *Rom J Intern Med* 53, 25–30. doi: 10.1515/rjim-2015-0003
- Mousavi Asl, E., Mahaki, B., Khanjani, S., and Mohammadian, Y. (2020). The Assessment of Alexithymia Across Positive and Negative Emotions: The Psychometric Properties of the Iranian Version of the Perth Alexithymia Questionnaire. *Iran J Psychiatry Behav Sci* 14, e102317. doi: 10.5812/ijpbs.102317
- Pandey, R., Saxena, P., and Dubey, A. (2011). Emotion regulation difficulties in alexithymia and mental health. *Eur J Psychol* 7, 604–623.
- Passardi, S., Peyk, P., Rufer, M., Wingenbach, T. S., and Pfaltz, M. C. (2019). Facial mimicry, facial emotion recognition and alexithymia in post-traumatic stress disorder. *Behav Res Ther* 122, 103436. doi: 10.1016/j.brat.2019.103436
- Preece, D. A., Becerra, R., Boyes, M. E., Northcott, C., McGillivray, L., and Hasking, P. A. (2020). Do self-report measures of alexithymia measure alexithymia or general psychological distress? A factor analytic examination across five samples. *Pers Individ Dif* 155, 109721. doi: 10.1016/j.paid.2019.109721
- Preece, D., Becerra, R., Allan, A., Robinson, K., and Dandy, J. (2017). Establishing the theoretical components of alexithymia via factor analysis: Introduction and validation of the attention-appraisal model of alexithymia. *Pers Individ Dif* 119, 341–352. doi: 10.1016/j.paid.2017.08.003
- Preece, D., Becerra, R., Robinson, K., and Dandy, J. (2018a). Assessing alexithymia: psychometric properties and factorial invariance of the 20-item Toronto alexithymia scale in nonclinical and psychiatric samples. *J Psychopathol Behav Assess* 40, 276–287. doi: 10.1007/s10862-017-9634-6
- Preece, D., Becerra, R., Robinson, K., Dandy, J., and Allan, A. (2018b). The psychometric assessment of alexithymia: Development and validation of the Perth Alexithymia Questionnaire. *Pers Individ Dif* 132, 32–44. doi: 10.1016/j.paid.2018.05.011
- Rudell, K., and Diefenbach, M. A. (2008). Current issues and new directions in psychology and health: culture and health psychology. Why health psychologists should care about culture. *Psychol Heal* 23, 387–390. doi: 10.1080/08870440701864983
- Saarijärvi, S., Salminen, J., and Toikka, T. (2001). Alexithymia and depression: a 1-year follow-up study in outpatients with major depression. *J Psychosom Res* 51, 729–733.
- Säkkinen, P., Kaltiala-Heino, R., Ranta, K., Haataja, R., and Joukamaa, M. (2007). Psychometric properties of the 20-item Toronto Alexithymia Scale and prevalence of alexithymia in a Finnish adolescent population. *Psychosomatics* 48, 154–161. doi: 10.1176/appi.psy.48.2.154
- Schumacker, R. E., and Lomax, R. G. (2016). *A beginner's guide to structural equation modeling*. New York, NY: Routledge.
- Sifneos, P. E. (1973). The prevalence of 'alexithymic' characteristics in psychosomatic patients. *Psychother Psychosom* 22, 255–262. doi: 10.1159/000286529
- Swart, M., Kortekaas, R., and Aleman, A. (2009). Dealing with feelings: characterization of trait alexithymia on emotion regulation strategies and cognitive-emotional processing. *PLoS one* 4:e5751. doi: 10.1371/journal.pone.0005751
- Taylor, G. J., Bagby, R. M., and Parker, J. D. (2003). The 20-Item Toronto Alexithymia Scale: IV. Reliability and factorial validity in different languages and cultures. *J Psychosom Res* 55, 277–283.
- Vorst, H. C., and Bermond, B. (2001). Validity and reliability of the Bermond-Vorst alexithymia questionnaire. *Pers Individ Dif* 30, 413–434. doi: 10.1016/s0191-8869(00)00033-7
- Wang, Y. P., and Gorenstein, C. (2013). Psychometric properties of the Beck Depression Inventory-II: a comprehensive review. *Braz J Psychiatry* 35, 416–431. doi: 10.1590/1516-4446-2012-1048
- Watters, C. A., Taylor, G. J., Quilty, L. C., and Bagby, R. M. (2016). An examination of the topology and measurement of the alexithymia construct using network analysis. *J Per Assess* 98, 649–659. doi: 10.1080/00223891.2016.1172077

**Conflict of Interest:** The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

Copyright © 2021 Lashkari, Dehghani, Sadeghi-Firoozabadi, Heidari and Khatibi. This is an open-access article distributed under the terms of the Creative Commons Attribution License (CC BY). The use, distribution or reproduction in other forums is permitted, provided the original author(s) and the copyright owner(s) are credited and that the original publication in this journal is cited, in accordance with accepted academic practice. No use, distribution or reproduction is permitted which does not comply with these terms.