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Differential temporal beta-diversity patterns of native and non-native arthropod species in a fragmented native forest landscape

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- 1 Journal: Ecography R2
- 2 Differential temporal beta-diversity patterns of native and non-native arthropod species
- 3 in a fragmented native forest landscape

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- An important factor that hinders the management of non-native species is a general lack of
- information regarding the biogeography of non-natives, and, in particular, their rates of
- 19 turnover. Here, we address this research gap by analysing differences in temporal beta-
- 20 diversity (using both pairwise and multiple-time dissimilarity metrics) between native and
- 21 non-native species, using a novel time-series dataset of arthropods sampled in native forest
- fragments in the Azores. We use a null model approach to determine whether temporal beta-
- 23 diversity was due to deterministic processes or stochastic colonisation and extinction events,
- 24 and linear modelling selection to assess the factors driving variation in temporal beta-
- 25 diversity between plots. In accordance with our predictions, we found that the temporal beta-
- 26 diversity was much greater for non-native species than for native species, and the null model
- 27 analyses indicated that the turnover of non-native species was due to stochastic events. No
- predictor variables were found to explain the turnover of native or non-native species. We
- 29 attribute the greater turnover of non-native species to source-sink processes and the close
- 30 proximity of anthropogenic habitats to the fragmented native forest plots sampled in our
- 31 study. Thus, our findings point to ways in which the study of turnover can be adapted for
- 32 future applications in habitat island systems. The implications of this for biodiversity
- conservation and management are significant. The high rate of stochastic turnover of non-
- native species indicates that attempts to simply reduce the populations of non-native species

- in situ within native habitats may not be successful. A more efficient management strategy
- would be to interrupt source-sink dynamics by improving the harsh boundaries between
- and adjacent anthropogenic habitats.
- 38 **Keywords:** Invasive species, beta-diversity, turnover, theory of island biogeography for
- 39 exotic species, island biogeography, fragmentation, habitat island

Introduction

- 41 The introduction, spread and establishment of species outside of their native range can result
- 42 in substantial changes to natural ecosystems (Mooney and Hobbs 2000, Dyer et al. 2017),
- 43 sometimes including the local and/or regional extirpation of native species (Clavero and
- García-Berthou 2005, Bellard et al. 2016). Furthermore, the global rate of transfer and mixing
- of species between native and non-native ranges does not show any signs of decreasing
- 46 (Seebens et al. 2017). We use the term 'non-native' throughout this study and we define such
- 47 species simply as those that are present in an area outside of their native range as a result of
- 48 human actions (Blackburn et al. 2016), which on islands is generally a consequence of
- 49 commerce, gardening, agriculture and forestry (Whittaker and Fernández-Palacios 2007).
- Indeed, the impact of non-native species, and particularly invasive species (a subset of non-
- native species), is variable but it does appear to be more acute on islands (Blackburn et al.
- 52 2004, Whittaker and Fernández-Palacios 2007, Bellard et al. 2016, but see Sax et al. 2002). A
- key component of assessing the impact of non-native species in island systems involves the
- development, and testing, of (island) biogeographical theories and models relevant to them
- 55 (Pyšek 1998, Blackburn et al. 2008, 2016, Burns 2015). For example, in their fifty
- 56 'fundamental questions in island biology', Patiño et al. (2017) recently highlighted that
- 57 understanding how the impacts of non-native species differ from those of naturally colonising
- species is a key question in their management.
- 59 Burns (2015) has started this process via a recently introduced extension of MacArthur and
- Wilson's (1967) equilibrium theory of island biogeography (ETIB), which he termed a
- 61 'theory of island biogeography for exotic species' (herein, 'TIBE'). TIBE is a graphical
- 62 island biogeographic model that makes a variety of different predictions regarding the species
- richness and turnover of native and non-native species (Burns 2015). This was a useful
- advance as little is known about the turnover patterns of non-natives. However, the analysis
- of turnover dynamics of non-natives has so far only been studied in the context of true

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islands, i.e. islands surrounded by a matrix of water (Whittaker and Fernández-Palacios 2007,
      Burns 2015). In contrast, the knowledge of turnover patterns for non-native species in habitat
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      islands, i.e. patches of natural habitat surrounded by a matrix of often human dominated
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      habitats (see Matthews 2015), is lacking. Such knowledge is important as the destruction and
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      fragmentation of natural habitat is widely recognised as the leading driver of contemporary
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      species extinctions (e.g. Sala et al. 2000) and also as an important driver of extinction debt
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      (Triantis et al. 2010). Moreover, there has been an increasing recognition of the synergistic
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      effects of the different drivers of species loss (e.g. habitat loss, climate change and invasive
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      species; e.g. Didham et al. 2007, Ferger et al. 2017, Karp et al. 2018). As with true islands,
      we predict that the turnover of non-native species will be greater than for native species in
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      habitat fragments. However, in true islands the matrix (water) can generally be ignored,
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      whilst in habitat island systems it is possible that the surrounding matrix contributes to
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      turnover patterns within habitat islands. As many non-native species have strong affinities to
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      human-dominated habitats, that is, they are generalists or human habitat specialists
      (McKinney and Lockwood 1999, Borges et al. 2008, 2010), the presence of non-native
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      species within habitat islands of native habitat is theorised to be driven by stochastic source-
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      sink mass effect dynamics as a result of human disturbance (Williamson 1996, Sgarbi and
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      Melo 2017). Non-native species, which should therefore be less adapted to the conditions
      within the sink habitat, should have a higher risk of extinction and turnover (MacArthur and
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      Wilson 1967). This possibility has not previously been examined in habitat islands.
      A variety of methods have been employed to analyse turnover in the island literature (e.g.
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      Russell et al. 1995, Whittaker et al. 2000, Burns 2015), mostly based on the summation of
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      raw numbers (e.g. number of extinction events). However, another, and arguably more
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      statistically robust, way of analysing temporal changes in species assemblages uses
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      dissimilarity indices, which allow researchers to partition out the effect of richness
      differences between samples (Baselga 2010, Baselga et al. 2015). One such approach is to
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      use the framework of temporal beta-diversity. Beta-diversity provides a measure of the
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      differences in the composition of communities, and is usually calculated in a spatial context,
      e.g. to assess how composition changes across a set of sites or along an ecological gradient
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      (Anderson et al. 2011). Temporal beta-diversity is a similar concept, where beta-diversity is
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      calculated for the same location at different times, and in conjunction with a suitable null
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      model the analysis of temporal beta-diversity can be used to determine whether changes in
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      assemblages across time are due to deterministic processes or stochastic colonisation and
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| 99 | extinction events (Baselga et al. 2015). Temporal beta-diversity sensu stricto has been much |
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| 100 | less studied relative to spatial beta-diversity (but see Baselga et al. 2015, Tonkin et al. 2017). |
| 101 | Beta-diversity can be calculated using a variety of different approaches (see Anderson et al. |
| 102 | 2011) and in this study we focus on the use of dissimilarity indices to calculate beta-diversity, |
| 103 | in particular Sørensen dissimilarity. Recent work has partitioned the Sørensen index into |
| 104 | turnover and nestedness-resultant dissimilarity / richness difference components (Baselga |
| 105 | 2010, 2012, Carvalho et al. 2012). |
| 106 | In this study, we use a unique time-series dataset of arthropods sampled in native forest |
| 107 | fragments over five years in the Azores (see Borges et al. 2017) to investigate the differences |
| 108 | in turnover dynamics of native and non-native species. Using TIBE and past studies on island |
| 109 | theory in habitat islands (e.g. Matthews 2015) as theoretical frameworks, we make two |
| 110 | predictions. First, based on the above points, we predict that temporal beta-diversity will be |
| 111 | greater for non-native species than for native species (Prediction 1). We use a null model |
| 112 | approach to determine whether turnover of species through time was due to deterministic |
| 113 | processes or stochastic colonisation and extinction events, and linear modelling selection to |
| 114 | assess if any factors (i.e. elevation, climate, disturbance) are driving variation in temporal |
| 115 | beta-diversity between plots. Second, as the invasion process is predicted to be on-going (see |
| 116 | Burns 2015, Seebens et al. 2017), based on the TIBE we predict that colonisation rate will be |
| 117 | greater than extinction rate for non-native species, whilst colonisation rate will be roughly |
| 118 | equivalent to extinction rate for native species (Prediction 2). |
| 119 | Our dataset is ideal for examining temporal beta-diversity patterns of native and non-native |
| 120 | species in habitat islands as the regular census interval allows us to accurately describe |
| 121 | colonisation and extinction events (and thus turnover), and the Azorean arthropod fauna |
| 122 | contains a high proportion of non-native species (Borges et al. 2010). Confirming or |
| 123 | invalidating these predictions will enable a better understanding of the turnover dynamics of |
| 124 | non-native species in fragmented landscapes and will provide important information to aid in |
| 125 | the conservation of fragmented natural areas impacted by the spread and establishment of |
| 126 | non-native species in currently threatened ecosystems. |

Material and methods

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Study site and data collection

| 129 | Arthropods were sampled using nine 50m x 50m plots located in four native fragments of |
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| 130 | pristine forest on Terceira Island in the Azores. The plots were setup within the |
| 131 | ISLANBIODIV project (Borges et al. 2017; Cicconardi et al. 2017). Arthropods were |
| 132 | sampled using a passive flight interception trap called a SLAM (Sea, Land, and Air Malaise) |
| 133 | trap. The collecting bottles were collected and changed every three months; thus, each sample |
| 134 | covers one season of the year. For the current study, we used data sampled over the years |
| 135 | 2012 -2016 (inclusive). The arthropods were grouped by their native and non-native |
| 136 | colonisation strategies. A more comprehensive outline of the study site (including a map) and |
| 137 | the sampling methodology is provided in Appendix S1 in the Supporting Information (see |
| 138 | also Borges et al. 2017). |
| 139 | For the climatic data, we used data from the CIELO Model (Azevedo et al. 1999). The |
| 140 | CIELO model is a simple parcel model, based on the transformations experienced by an air |
| 141 | mass ascending a mountain, which simulates the evolution of an air parcel's physical |
| 142 | properties, starting from the sea level. Two Principal Components Analysis (PCA) axes were |
| 143 | calculated using the climatic variables mean annual temperature, annual rainfall and relative |
| 144 | humidity. The PCA was undertaken using the 'vegan' R package (Oksanen et al. 2013), and |
| 145 | we took the first two axes as they explained ~ 99% of the variance. The first axis (P1) |
| 146 | corresponded to differences in temperature and precipitation, whilst the second axis (P2) |
| 147 | corresponded more to differences in humidity. In addition, we used an 'index of Disturbance' |
| 148 | developed by Cardoso et al. (2013) for the Azores that models disturbance by considering |
| 149 | landscape configuration and proximity of human-modified habitat to each patch of pristine |
| 150 | native forest (see Appendix S1). We also calculated the elevation of each plot using a digital |
| 151 | elevation model (DEM) for Terceira Island. |
| 152 | We pooled the samples within each year to create yearly datasets for each of the nine plots. |
| 153 | Where the same species had been sampled in multiple samples within a single year we |
| 154 | combined records and summed the abundances. We then converted these data into presence- |
| 155 | absence matrices, such that for each of the nine plots we had five presence-absence matrices, |
| 156 | one for each of the five sampling years (2012–2016). |
| 157 | Calculating temporal beta-diversity: Prediction 1 |
| 158 | To examine whether temporal beta-diversity was greater for non-natives than for natives |
| 159 | (Prediction 1), we used both pairwise and multiple-time temporal beta-diversity frameworks. |

First, for each plot we constructed a presence-absence matrix for the pooled 2013 samples

161 (the first full year of sampling, see Appendix S1) and the pooled 2016 samples. We then used the 'beta.temp' function in the 'betapart' R package (Baselga et al. 2017) to calculate 162 partitioned pairwise temporal beta-diversity (Sorensen index) between these two years, and 163 we stored both partition values (i.e. temporal turnover and temporal nestedness-dissimilarity) 164 as well as the overall Sorensen's dissimilarity value. This analysis was undertaken separately 165 for native and non-native species. To determine whether the observed pairwise temporal beta-166 diversity and partition values were significantly different from random expectation, we 167 followed Baselga et al. (2015) and used an FE null model (Ulrich and Gotelli 2007) whereby 168 169 row (site species richness) marginal totals were allowed to vary and column (species incidences) marginal totals are fixed. In this context, the FE null model characterises a 170 situation whereby species randomly colonise and go extinct at sites, from a common regional 171 pool (Baselga et al. 2015). This null model was implemented using the 'c0' algorithm in the 172 'vegan' R package (Oksanen et al. 2013). This null model was used to randomise the 2016 173 presence/absence matrix, which was then compared to the empirical 2013 presence/absence 174 matrix, using the 'beta.temp' function. This process was repeated 1000 times, for both the 175 native and the non-native species data; the overall dissimilarity value, and the two partition 176 values, was stored in each instance. Using these null model values, we calculated the Z-177 178 scores for the six metrics (the three temporal beta-diversity metrics, for both the native and non-native presence/absence matrices). A two-sided P-value was also calculated for each Z-179 180 score. To ensure that our results were consistent across the two main beta-diversity partitioning frameworks, we also calculated temporal beta-diversity and re-ran the above 181 182 analyses using the Carvalho partitioning framework (Carvalho et al. 2012), the Sorensen family of beta-diversity metrics and the BAT R package (Cardoso et al. 2014). The Carvalho 183 184 partitioning framework differs from the Baselga framework in that it partitions overall betadiversity into turnover and richness difference components, instead of turnover and 185 nestedness-dissimilarity components. 186 Spatial beta-diversity studies have shown that pairwise beta-diversity metrics and multisite 187 188 beta-diversity metrics (i.e. spatial beta-diversity aggregated across multiple sites) can reveal different patterns (Baselga 2013). Thus, we adapted the 'beta.temp' function to calculate 189 partitioned multiple-time dissimilarity, again based on the Sorensen index. This is an 190 aggregate measure that enables the calculation of temporal beta-diversity across multiple time 191 periods for the same site. We calculated multiple-time dissimilarity for both native and non-192

193 native species for four years (2013 - 2016) and compared the resultant beta-diversity values with their pairwise counterparts. 194 **Calculating turnover: Prediction 2** 195 196 Following Burns (2015), turnover was calculated as the total number of colonisation and extinction events across the five years (2012 – 2016). That is, colonisation was considered to 197 198 have occurred if a species was not present in a plot in year i, but was present in year i+1. Equally, under this turnover framework, an extinction event was considered to have occurred 199 200 if a species was present in a plot in year i, but not in year i+1. To test Prediction 2 (that colonisation rate will be greater than extinction rate for non-native 201 202 species, but equivalent to extinction rate for native species), we calculated the number of 203 turnover events for each individual species, across all nine plots. Reduced major axis regression was then used to determine how colonisation rate varied with extinction rate (see 204 Burns 2015) as both variables (colonisation and extinction rate) contained random error; in 205 such cases, standard linear regression underestimates the slope of the relationship. The 206 regression models were calculated using the 'Imodel2' R package (Legendre 2014), and we 207 used the standard major axis (SMA) method. Models were fitted for native and non-native 208 species separately. In this analysis, a slope significantly greater than one, or an intercept 209 210 significantly greater than zero, would indicate that the colonisation rate exceeded extinction rate (for either native or non-native species) and thus that the number of native/non-native 211 212 species increased in the fragments over the five years, and *vice versa*. 213 **Explaining variation in temporal beta-diversity** To determine whether any of our environmental variables explained variation in the turnover 214 215 component of temporal beta-diversity, we undertook a model comparison approach using 216 generalised linear models (GLMs). We used GLMs with the Gaussian family in these analyses, and normality of the response variables was assessed using the Shapiro-Wilks test 217 218 (in both cases the null hypothesis that the response variable was normally distributed could not be rejected). We ran the model comparison twice, once for each of two response 219 variables: 1) the pairwise temporal turnover beta-diversity partition values of native species, 220 and 2) the pairwise temporal turnover beta-diversity partition values of non-native species. 221 As predictor variables, we started with elevation (log transformed), disturbance (log 222 transformed) and two climatic PCA axes (P1 and P2). All predictors were standardised to 223

have a mean of 0 and standard deviation of 1 to enable comparison of the effect sizes.

| 225 | Multiconflicantly was assessed using variance infration factors (virs), which resulted in F1 |
|-----|---|
| 226 | being removed from all subsequent analyses as it was very highly correlated with altitude. |
| 227 | The three remaining predictor variables (elevation, disturbance and P2) all had VIFs under |
| 228 | ten. |
| 229 | Using these response and predictor variables, we compared a full set of generalised linear |
| 230 | models (GLMs) within an information theoretic approach (Burnham and Anderson 2002). |
| 231 | For each response variable, a full set of models were compared using Akaike's information |
| 232 | criterion corrected for small sample size (AICc; see Burnham and Anderson 2002). The |
| 233 | model with the lowest AIC $_{c}$ was considered the best model, whilst all models with $\Delta AIC_{c}\!<=\!$ |
| 234 | 2 were considered to have similar support. As our data were sampled using 50 m x 50 m |
| 235 | quadrats nested within fragments, we used a spatial autocovariate within an auto-Gaussian |
| 236 | regression approach to account for the experimental design. First, for each response variable, |
| 237 | we fitted the global model and extracted the residuals. We then created a spatial autocovariate |
| 238 | using the residuals (see Crase et al. 2012) and the 'autocov_dist' function in the 'spdep' R |
| 239 | package (Bivand 2017). The neighbourhood radius was set to encompass all plots, and we |
| 240 | used the 'inverse' type and row standardised (W) style settings. The spatial autocovariate was |
| 241 | fixed in the model selection. For all models with $\Delta AIC_c \ll 2$, we also checked the residuals |
| 242 | for homoscedasticity, and we checked for any remaining spatial autocorrelation in the |
| 243 | residuals using the 'spdep' R package (Bivand 2017), the 'nb2listw' function and row |
| 244 | standardised weights. We also re-ran the model selection analyses using Gaussian mixed |
| 245 | effect models with 'fragment' as a random effect (random intercept); however, the variance |
| 246 | of the random effect was very close to zero in both cases and these results are not discussed |
| 247 | further. |
| 248 | Sensitivity analyses |
| 249 | We ran two sensitivity analyses to ensure our results were robust. First, to ensure our |
| 250 | sampling was sufficient we calculated sampling completeness estimates for each year in each |
| 251 | plot using the iNEXT R package (Hill number order $q = 0$, Hsieh et al. 2016). Second, we set |
| 252 | up a tenth plot in which we placed three SLAM traps instead of one to determine whether the |
| 253 | use of a single SLAM trap in each of the plots was sufficient to capture the relevant |
| 254 | community properties. A full methodology for each of the sensitivity analyses is provided in |

Appendix S1. All analyses were undertaken using R (version 3.4.3, R Core Team 2017).

Results

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Dataset summary

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Over the five years, we sampled 28,704 arthropod specimens, representing 147 species and morpho species (no. of native sp. = 89 and no. of non-native sp. = 58), across ten plots (nine plots were used in the main analysis and a tenth plot was used in the sensitivity tests). Across the nine plots that formed the main basis of this study, the mean richness (i.e. mean of each plot across the five years) of species in a plot ranged from 21 to 38 and 3 to 11, for native and non-native species respectively (Table 1). Mean total abundance (i.e. abundance of all species in a plot) ranged from 205 to 886 and 4 to 68 for native and non-native species respectively (Table 1).

Temporal beta-diversity

For all nine plots, overall temporal beta-diversity was larger for non-native species than for 267 native species (Table 2; Fig. 1), confirming Prediction 1. The temporal turnover component 268 of overall pairwise beta-diversity was larger than the temporal nestedness-dissimilarity 269 component for all but one plot for non-native species, and for five of the nine plots for native 270 species (Fig. 1). For overall temporal beta-diversity, the Z-scores were negative for all but 271 one plot for native species, whilst Z-scores were both positive and negative for non-native 272 species (Table 2). Four of the overall pairwise beta Z-scores for native species were 273 significant, whilst only one of the overall pairwise beta Z-scores for non-native species was 274 significant (Table 2). The null model results for the pairwise temporal beta-diversity 275 partitions were similar to the results for the overall pairwise temporal beta-diversity values 276 277 (see Table 2). This outcome did not change when the Carvalho pairwise beta-diversity partitioning framework was used rather than the Baselga approach (see Appendix S2): overall 278 279 pairwise beta-diversity and the two partitions were all positively and significantly correlated between the two approaches (all Pearson's correlation coefficients > 0.90, and all P-values < 280 281 0.001), and the same overall picture emerged regardless of the approach used (compare Fig. 1 282 with Fig. S3 in Appendix S2). Using multiple-time dissimilarity produced similar results to the pairwise temporal beta-283 diversity analysis (see Figure S4 in Appendix S3): overall temporal beta-diversity was larger 284 for non-native species than for native species, for all nine plots, and on average the temporal 285 nestedness-dissimilarity component represented a larger proportion of total temporal beta-286 diversity for native species than for non-native species. One interesting difference was non-287

native species in Plot 6 (compare Fig.1 with Fig.S4), whereby the pairwise measure indicated

| 289 | that the nestedness-dissimilarity component represented 100% of overall beta-diversity, but |
|-----|---|
| 290 | only 16% using multiple-time dissimilarity. |
| 291 | Differences in colonisation and extinction rates |
| 292 | Reduced major axis regression of the number of colonisation events against the number of |
| 293 | extinction events revealed that, contrary to Prediction 2, the slope of the relationship was |
| 294 | significantly greater than one for both native (slope = 1.26 ; 95% CI = $1.11 - 1.44$) and non- |
| 295 | native species (slope = 1.28 ; 95% CI = $1.16 - 1.42$) (Fig. 2), indicating that the colonisation |
| 296 | rate exceeded the extinction rate for both groups. The intercept of the relationship was not |
| 297 | significantly different from zero for either native (intercept = 0.17 ; 95% CI = $-0.22 - 0.51$) or |
| 298 | non-native (intercept = -0.11 ; 95% CI = $-0.38 - 0.13$) species. The R ² values for both |
| 299 | regression models were: 0.64 and 0.86 for the native and non-native species models, |
| 300 | respectively. |
| 301 | Explaining variation in temporal beta-diversity between plots |
| 302 | When the turnover partition (of overall pairwise temporal beta-diversity) values were used as |
| 303 | the response variables in a set of GLMs, the selection procedure indicated that the only |
| 304 | predictor variable in both the best native and non-native species models was the fixed spatial |
| 305 | autocovariate, and there were no additional models within 2 $\Delta AIC_c of$ either best model. Re- |
| 306 | running the model selection using the turnover component of multiple-time dissimilarity |
| 307 | produced the same overall results. |
| 308 | Sensitivity analyses |
| 309 | Sample completeness estimates indicated that our sampling was sufficient: the mean sample |
| 310 | completeness estimate across all years and plots was 0.97 (range = 0.81 to 1.00 , the results |
| 311 | for each year and plot are presented in Table S2 in Appendix S4). Using three SLAM traps in |
| 312 | a plot rather than one did not result in substantially different sampled communities according |
| 313 | to various measures of diversity (see Appendix S4). |
| 314 | |
| 315 | Discussion |
| 316 | We have undertaken an evaluation of the differences in temporal beta-diversity patterns |
| 317 | between native and non-native arthropod species across nine native forest plots in the Azores |
| 318 | for which a time series of five years is available (2012 - 2016). |

Azorean forest fragments 320 The results of our temporal beta-diversity analyses (both pairwise and multiple-time 321 322 dissimilarity) indicated that, in accordance with Prediction 1, temporal beta-diversity was considerably greater for non-native species than for native species, across all nine plots. In 323 addition, with one exception (Plot 6 when using pairwise beta-diversity, Fig. 1), for non-324 native species the largest component of temporal beta-diversity was by far the temporal 325 turnover partition, indicating an actual turnover of species rather than nested patterns in 326 richness (Baselga partition) or richness differences (Carvalho partition) between years. 327 328 Burns (2015), in his theory of island biogeography for exotic species (TIBE), attributed differences in turnover rates between natives and non-natives to standard island 329 330 biogeographical variables. However, Burns was focused on true islands. In our study system of habitat islands, we found that, in regard to overall pairwise temporal beta-diversity for 331 non-native species, the null model analyses were only significant in one case, and the signs of 332 the Z-scores varied between the plots (Table 2). These findings, in combination with previous 333 work in our study system (e.g. Borges et al. 2006, 2008, Florencio et al. 2016), appear to 334 support the notion that the distribution of non-native arthropods in Azorean native forests is 335 driven by stochastic processes and occupancy dynamics, possibly due to source-sink 336 processes (see also Borges et al. 2008). In a study of temporal beta-diversity patterns of birds 337 in France, Baselga et al. (2015, p.9) also found that temporal changes in assemblages were 338 not significantly different from a null model and concluded that "the observed changes in 339 340 species composition of local bird assemblages might be the consequence of stochastic 341 processes in which species populations appeared and disappeared from specific localities in a random-like way." 342 343 Many non-native arthropod species in the Azores are adapted to human land-uses (Borges et al. 2008, Rigal et al. 2018), and thus we would expect there be a large number, and thus high 344 345 potential for mass effects (Shmida and Wilson 1985), of non-natives in the disturbed landscapes surrounding the fragments (see also Borges et al. 2006). The Azores has 346 347 undergone substantial land use change since human colonisation of the archipelago (Triantis et al. 2010), and native forest fragments in the Azores are characterised by hard boundaries, 348 i.e. there is an abrupt change from native forest habitat to anthropogenic habitat (Borges et al. 349 2006, 2008). As such, there are likely to be large mass effects and a constant supply of non-350

Non-native species have larger rates of temporal beta-diversity than native species in

native individuals permeating into the native forest where they frequently undergo local extinction and re-colonisation, leading to high beta-diversity through time. Thus, it seems likely that is not just the size of the native habitat that underpins the colonisation rate of non-natives (as in true islands cf. Burns 2015) but also the amount of surrounding anthropogenic habitat and size of the non-native source pool, which is known to be large in the Azores (Borges et al. 2010). More detailed studies focusing on the habitat affinities and dispersal ecology of non-native species are needed to further explore this possibility. Interestingly, our disturbance metric, that incorporates surrounding land use, was not an important predictor of non-native temporal beta-diversity in the linear model selection analysis. However, as all of the native forest fragments on Terceira are surrounded by human land uses, it is likely that there was simply not enough variation in the disturbance metric between plots.

In regards to the processes underpinning temporal beta-diversity patterns of native species, the results are more equivocal. Overall pairwise temporal beta-diversity was significantly lower than expected by chance in four of the nine plots for native species, in comparison to only one of the nine plots for non-native species. This indicates a stronger role for deterministic processes driving temporal beta-diversity and turnover in native species assemblages (Baselga et al. 2015). However, as with non-native species, none of our predictor variables were found to explain variation in the temporal beta-diversity of native species. The reasons for this finding are unclear but could simply be due to the fact that our experimental design did not allow us to test for the importance of other biogeographic variables such as area on turnover (MacArthur & Wilson 1967). Further studies examining the turnover dynamics of native and non-native species in habitat fragments are needed.

Colonization and extinction rates for native and non-native species

Our Prediction 2, that colonisation rate will be greater than extinction rate for non-native species, whilst colonisation rate will be roughly equivalent to extinction rate for native species, was not borne out by the data. Whilst we observed that the colonisation rate was greater than the extinction rate for non-natives, contrary to expectations we found a similar pattern for native species (Fig. 2). Whilst this finding is expected for non-native species, these results could imply either that many of the native species in the fragments have high dispersal ability and are easily able to disperse between plots and fragments, or possibly that native species have not yet reached equilibrium following substantial habitat loss over the last few centuries. An alternative explanation is that our sampling did not accurately record all

colonisation and extinction events during the study period. Previous work on turnover on islands has shown that the calculation of turnover rate is sensitive to the grain size of the time series data analysed (e.g. were samples carried out every year or every ten years) (e.g. Russell et al. 1995, Whittaker et al. 2000). Considering these issues, inevitably we have not recorded every true turnover event (i.e. crypto-turnover), and equally, due to sampling error, we have likely missed individuals in certain instances and thus erroneously recorded turnover events (i.e. pseudo-turnover). However, sampling was every three months (as opposed to multiple years in most island studies), and thus a species had to be absent across all the monthly samples for an extinction event to be classified. We also carried out sensitivity analyses to ensure our sampling effort was sufficient. As a result, we are confident that our sampling protocol has generated data of sufficient quality to test our predictions.

Implications for conservation and biodiversity management

Non-native species are thought to be leading drivers of contemporary species extinctions, and the issue seems particularly acute in true and habitat island systems (Sax and Gaines 2003, Cardoso et al. 2010, Bellard et al. 2016). However, the biogeography of non-native species is not well known. Specifically, it is not known whether standard biogeographical theory and metrics derived from the study of native species can be accurately applied to non-native taxa. As a result of this uncertainty, a number of recent studies have focused on examining the differences and similarities between natives and non-natives using classic biogeographical and ecological patterns, such as the ISAR and the abundance-occupancy relationship (e.g. Blackburn et al. 2008, Rigal et al. 2013, Burns 2015). The results of our study contribute towards filling this knowledge gap by showing that it cannot be assumed *a priori* that native and non-native taxa within the same community are similarly assembled and will follow the same temporal dynamics.

Importantly, our results indicate that it might not be possible to apply simple biogeographical principles (e.g. Wilson and Willis 1975) when devising non-native species management plans in fragmented landscapes. Rather, management should be based on the results of biogeographical and ecological studies explicitly focused on non-native taxa. A corollary of this statement is that there is an urgent need for additional studies focused on outlining and testing biogeographical theory in the context of non-native species (Patiño et al. 2017). In the native forest sampled in our study, the high rate of stochastic turnover of non-native species that we observed indicates that attempts to simply reduce the populations of non-native

| species <i>in situ</i> within native habitats may not be successful. This is because for many non- |
|--|
| native species the native habitat is likely just a sink, and the constant immigration of |
| individuals from anthropogenic source habitats in close proximity to the native habitats (i.e. |
| mass effects) means local extinction of non-natives within the native forest is unlikely to be |
| permanent. Thus, a more efficient future management strategy and land-use policy will be to |
| interrupt these source-sink dynamics by improving the harsh boundaries between native |
| habitat and adjacent anthropogenic habitat, and in the longer term to design and create a more |
| graded landscape mosaic (Lindenmayer and Fischer 2006) whereby contrasting land uses are |
| not simply knitted together in an ad hoc fashion. The high turnover of non-native species also |
| brings into question the ability of non-natives to, amongst other things, replace the functional |
| roles of extirpated native species (see Whittaker et al. 2014; Rigal et al. 2018) or increase |
| functional redundancy in fragmented landscapes. Functional diversity studies that focus on |
| non-native species are often based on static time periods (e.g. a sample from one year) and |
| their conclusions thus do not account for the fact that many non-native species present in a |
| patch may simply be ephemeral members of a given community. |

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Author contributions – TM designed the analytical study. PB designed the sampling

methodology and collected the data with RC and RN. TM analyzed the data. TM wrote the

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manuscript, with the help of JS and PB.

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Supporting Information

Additional Supporting Information may be found in the online version of this article.

Appendix S1. Additional methodological information, including plot locations and environmental data.

Appendix S2. Pairwise beta-diversity results using the Carvalho partition framework

Appendix S3. Multiple-time dissimilarity results

Appendix S4. Sensitivity Analyses Full Results

TABLES

Table 1. The mean richness (M. Richness), mean abundance (M. Abundance), and the number of colonisation (Colonis.), extinction, and total turnover events of arthropod species across nine native forest plots on Terceira Island, in the Azores. For each plot, the data are provided for native (Nat) and non-native (Non) species separately. Each plot was sampled multiple times across five years and samples were pooled to create five yearly samples (2012 – 2016). An extinction event was deemed to have occurred if a species was present in year i but not in year i+1, and *vice versa* for a colonisation event (Plot notation as in Table S1 and Figure S1).

| Plot | Type | M. Richness | M. Abundance | Colonis. | Extinction | Turnover |
|------|------|-------------|--------------|----------|------------|----------|
| 1 | Nat | 32.8 | 790.8 | 27 | 24 | 51 |
| 2 | Nat | 30.2 | 496.4 | 28 | 26 | 54 |
| 3 | Nat | 25.6 | 558.0 | 19 | 20 | 39 |
| 4 | Nat | 38.0 | 885.6 | 46 | 20 | 66 |
| 5 | Nat | 25.0 | 511.0 | 36 | 17 | 53 |
| 6 | Nat | 21.0 | 204.6 | 26 | 26 | 52 |
| 7 | Nat | 25.0 | 272.8 | 28 | 21 | 49 |
| 8 | Nat | 20.6 | 480.2 | 21 | 16 | 37 |
| 9 | Nat | 30.2 | 489.4 | 27 | 22 | 49 |
| 1 | Non | 6.8 | 20.8 | 17 | 10 | 27 |
| 2 | Non | 4.6 | 9.4 | 10 | 11 | 21 |
| 3 | Non | 4.0 | 5.6 | 10 | 9 | 19 |
| 4 | Non | 10.8 | 67.6 | 31 | 21 | 52 |
| 5 | Non | 4.25 | 11.2 | 10 | 8 | 18 |
| 6 | Non | 3.75 | 4.4 | 13 | 12 | 25 |
| 7 | Non | 3.2 | 5.0 | 8 | 7 | 15 |

8 Non 4.4 8.8 19 14 33 9 Non 5.0 18.2 13 15 28

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Table 2. Pairwise temporal beta-diversity values for arthropod species sampled in 2013 and 2016 in nine native forest plots on Terceira Island, in the Azores. For each plot, the data are provided for native (Nat) and non-native (Non) species separately. Overall temporal beta-diversity values (Sorensen dissimilarity index) are provided in addition to the temporal turnover (Turn.) and temporal nestedness-dissimilarity (Nest.) components of overall temporal beta-diversity. For the two partition values and the overall temporal beta-diversity value, significance was determined using an FE null model (1000 iterations). For each of the three beta-diversity values, the Z-score (Z) and associated *P*-value (*P*) are provided (see 'Materials and methods'). *P*-values significant at the 0.05 level are highlighted in bold (Plot notation as in Table S1 and Figure S1). The overall temporal beta-diversity differ very slightly from the sum of the turnover and nestedness components in certain plots due to rounding error.

| Plot | Type | Temporal beta-diversity | | | Turr | over | Nestedness | | Overall | |
|------|------|-------------------------|-------|---------|-------|--------|------------|--------|---------|--------|
| | | Turn. | Nest. | Overall | Z | P | Z | P | Z | P |
| 1 | Nat | 0.12 | 0.05 | 0.18 | -0.94 | 0.35 | -1.48 | 0.14 | -2.49 | 0.01 |
| 2 | Nat | 0.08 | 0.23 | 0.30 | -0.85 | 0.39 | 0.52 | 0.60 | -0.33 | 0.74 |
| 3 | Nat | 0.05 | 0.14 | 0.18 | -2.92 | < 0.01 | 3.64 | < 0.01 | -1.08 | 0.28 |
| 4 | Nat | 0.11 | 0.08 | 0.18 | -0.36 | 0.72 | -1.58 | 0.11 | -1.96 | 0.04 |
| 5 | Nat | 0.04 | 0.06 | 0.10 | -2.04 | 0.04 | -0.24 | 0.81 | -2.53 | 0.01 |
| 6 | Nat | 0.11 | 0.18 | 0.29 | -2.68 | 0.01 | 6.17 | < 0.01 | -0.1 | 0.92 |
| 7 | Nat | 0.12 | 0.11 | 0.23 | -1.00 | 0.32 | 0.75 | 0.45 | -0.44 | 0.66 |
| 8 | Nat | 0.15 | 0.08 | 0.23 | -3.45 | < 0.01 | 1.22 | 0.22 | -3.48 | < 0.01 |
| 9 | Nat | 0.17 | 0.09 | 0.26 | 0.11 | 0.91 | -0.72 | 0.47 | -0.59 | 0.56 |
| 1 | Non | 0.22 | 0.04 | 0.26 | -1.13 | 0.26 | -0.94 | 0.35 | -2.58 | 0.01 |
| 2 | Non | 0.75 | 0.05 | 0.80 | 1.50 | 0.13 | -0.39 | 0.70 | 1.63 | 0.10 |
| 3 | Non | 0.25 | 0.15 | 0.40 | -1.52 | 0.13 | 0.74 | 0.46 | -1.48 | 0.14 |
| 4 | Non | 0.44 | 0.12 | 0.57 | -0.24 | 0.81 | -0.27 | 0.79 | -0.63 | 0.53 |
| 5 | Non | 0.5 | 0.00 | 0.50 | 0.75 | 0.45 | -1.72 | 0.08 | -0.47 | 0.64 |
| 6 | Non | 0.00 | 0.71 | 0.71 | -4.84 | < 0.01 | 11.20 | < 0.01 | -0.72 | 0.47 |
| 7 | Non | 0.5 | 0.25 | 0.75 | -0.38 | 0.70 | 1.73 | 0.08 | 0.69 | 0.49 |
| 8 | Non | 0.8 | 0.03 | 0.83 | 0.02 | 0.98 | 0.06 | 0.95 | 0.05 | 0.96 |
| 9 | Non | 0.33 | 0.3 | 0.64 | -1.06 | 0.29 | 1.63 | 0.10 | -0.09 | 0.92 |

FIGURES

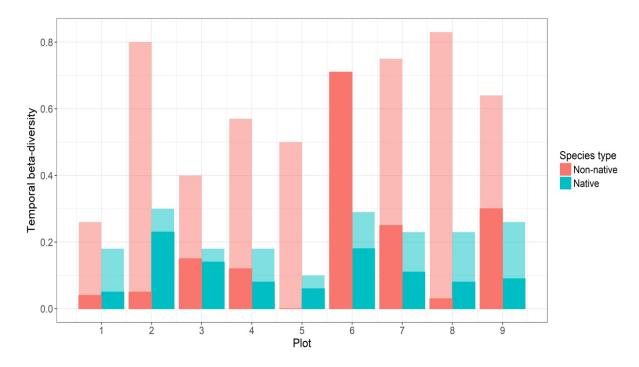


Figure 1. Pairwise temporal beta-diversity values for native and non-native arthropod species across nine plots of native forest on Terceira Island, in the Azores. For each plot, temporal beta-diversity was calculated using the pooled 2013 samples and the pooled 2016 samples. For each plot, temporal beta-diversity was calculated separately for native species (blue bars) and non-native species (red bars). The height of each bar corresponds to the overall temporal beta-diversity. Overall temporal beta was also partitioned into nestedness and turnover components using the Baselga partition approach (Baselga et al. 2017). Within each bar, the dark shaded area rising from the x-axis corresponds to the nestedness component of overall temporal beta diversity. Thus, the lighter shaded area within each bar that rises from the dark shaded area corresponds to the turnover component.

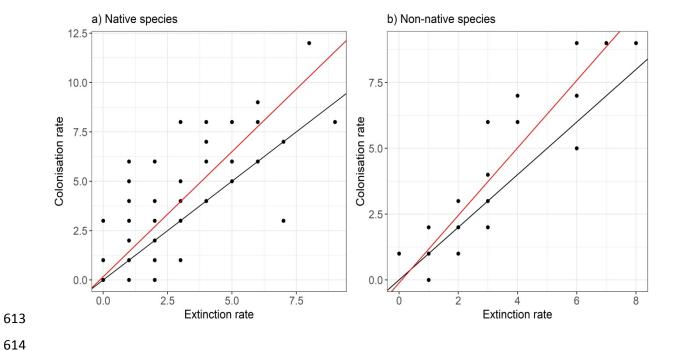


Figure 2. The relationship between the number of colonisation events and the number of extinction events in native (a) and non-native (b) arthropod species sampled in nine native forest plots on Terceira Island in the Azores. Each plot was sampled multiple times across five years and samples were pooled to create five yearly samples. Turnover was then calculated as the total number of colonisation and extinction events across the five years (see 'Materials and methods'). In (a) and (b), the black line is the isometric line (i.e. intercept of zero and slope of 1) and the red line is the best fit line from reduced major axis regression. In both cases the best fit regression line is significantly greater than one.

| 627 | SUPPORTING INFORMATION |
|------------|--|
| 628 629 | Differential temporal beta-diversity patterns of native and non-native arthropod species in a fragmented native forest landscape |
| 630 | Thomas J. Matthews, Jon Sadler, Rui Carvalho, Rui Nunes, Paulo A. V. Borges |
| 631 | |
| 632 | Appendix S1 Additional methodological information |
| 633 | Study site |
| 634 | The Azorean archipelago is located in the North Atlantic, roughly between 37° to 40° N |
| 635 | latitude and 25° to 3° W longitude. Since human occupation of the islands, there have been |
| 636 | substantial changes in the size and quality of native habitats, mostly due to the creation of |
| 637 | urban areas, agriculture fields, pastures and non-native plantations (Borges et al. 2008, |
| 638 | Triantis et al. 2010). Terceira Island (Fig. 1), a roughly circular island of 402 km ² area, |
| 639 | contains the largest area of native pristine forests in the Azores (Triantis et al. 2010), with |
| 640 | five main fragments of native forest distributed across four main volcanic polygenetic |
| 641 | complexes. |
| 642 | Data collection |
| 643 | Arthropods were sampled using nine 50m x 50m plots located in four native fragments of |
| 644 | pristine forest on Terceira setup within the ISLANBIODIV project (Cicconardi et al. 2017): |
| 645 | Serra de Santa Bárbara (Plots T07, T48, T49, T164), Biscoito da Ferraria (Plots T01, T02, |
| 646 | T41), Terra Brava (Plot T15) and Galhardo (Plot T33) (see Figure S1, below). Arthropods |
| 647 | were sampled using a passive flight interception trap called a SLAM (Sea, Land, and Air |
| 648 | Malaise) trap. The SLAM traps are approximately 110 x 110 x 110 cm (see Fig. S2 in |
| 649 | Appendix S1), and work by intercepting arthropods on an area of black mesh and funnelling |
| 650 | them into a sampling bottle filled with a killing liquid. Propylene-glycol was used as it |
| 651 | persists for a long time without evaporating, and enables the collection of good quality |
| 652 | specimens for posterior DNA extraction. The collecting bottles were collected and changed |
| 653 | every three months; thus, each sample covers one season of the year. However, due to |
| 654 | logistical challenges and problems with the traps, in a very small number of cases (see Table |
| 655 | S1 in Appendix S1) certain three-month samples were excluded. Further details on the |
| 656 | method can be found in Borges et al. (2017). Samples were sorted and individuals identified |
| 657 | to species level by experienced taxonomists and to morpho species where specific |
| 658 | determinations were not possible. The arthropods were then grouped by their native and non- |

659 native colonisation strategies. For the current study, we used data sampled over the years 2012 –2016 (inclusive). Sampling only commenced half way through 2012 (see Table S1, 660 below) and thus for the temporal beta-diversity analyses we used 2013 as the base year to 661 ensure that the pairwise comparisons were based on equal samples. However, for the species-662 level turnover analyses we used the full dataset as the missing months here are less likely to 663 affect the results. 664 For the climatic data, we used the CIELO Model (Azevedo et al. 1999). The CIELO model is 665 a simple parcel model, based on the transformations experienced by an air mass ascending a 666 mountain, which simulates the evolution of an air parcel's physical properties, starting from 667 the sea level. CIELO climatic data for the sites were extracted from Borges et al. (2006). In 668 addition, we used an 'index of Disturbance' developed by Cardoso et al. (2013) for the 669 Azores that models disturbance by considering landscape configuration and proximity of 670 human-modified habitat to each patch of pristine native forest. Thus, in this study we 671 represent anthropogenic disturbance of a plot by explicitly considering landscape 672 configuration and the amount of neighboring anthropogenic habitats, ranked according to 673 their level of intensity-use. 674 **Species sorting and classification** 675 Parataxonomists sorted samples to orders, and posteriorly to Recognizable Taxonomic Units 676 (RTUs). One of the authors (PAVB) then identified to species level the RTUs of the 677 following arthropod orders: Diplopoda (Chordeumatida, Julida), Chilopoda 678 679 (Geophilomorpha, Lithobiomorpha, Scolopendromorpha), Arachnida (Araneae, Opiliones, Pseudoscorpiones) and Insecta (Blattaria, Coleoptera, Hemiptera, Microcoryphia, Neuroptera, 680 681 Psocoptera, Thysanoptera, Trichoptera). All material is stored at EDTP—Entomoteca 682 Dalberto Teixeira Pombo, University of Azores, Angra do Heroísmo, Portugal. Arthropods were grouped into three colonization categories: endemic (i.e. restricted to the 683 684 Azores); native non-endemic, i.e. species that arrived naturally to the archipelago but are present both in the Azorean Islands and elsewhere; and non-native species, i.e., species 685 whose original distribution range did not include the Azores and that are believed to have 686 been introduced in the Macaronesian region after human settlement in the 15th century. The 687 688 non-native status was inferred either from historical records of detected species introductions 689 or from their current distribution being closely associated with human activity. For 690 unidentified species, if other species in the same genus, subfamily or family were present in

| 691 | the archipelago and all belonged to the same colonization category (according to Borges et al. |
|-----|---|
| 692 | 2010), the unknown species were classified similarly. Otherwise, we assumed the species to |
| 693 | be native. For simplicity, endemic and native non-endemic species were grouped and are |
| 694 | termed "native" throughout the text. |
| 695 | Sensitivity analyses methodology |
| 696 | To ensure our sampling was sufficient we calculated sampling completeness estimates for |
| 697 | each year in each plot using the iNEXT R package (Hill number order $\mathbf{q}=0$, Hsieh et al. |
| 698 | 2016). |
| 699 | Due to the resources required to sample arthropods in multiple plots every three months over |
| 700 | five years, we only placed one SLAM trap in each plot. To determine whether this was |
| 701 | sufficient to capture the relevant properties (e.g. species composition, distribution of |
| 702 | abundance) of the sampled communities, in 2015 we set up a tenth plot in Terra Brava (Plot |
| 703 | T18), in which we placed three SLAM traps, spaced 30 m apart, and sampled every month. |
| 704 | We constructed three overall assemblage matrices whereby, for each SLAM trap, the data |
| 705 | from the 12 months of 2015 were pooled into one sample. To determine whether community |
| 706 | composition was similar across the three traps, we calculated a number of summary statistics: |
| 707 | 1) the number of species in each assemblage, 2) the total abundance of each assemblage, 3) |
| 708 | the proportion of species that were sampled in all three traps (i.e. are present in all three |
| 709 | assemblages), and 4) the distribution of abundance across the species in each assemblage. In |
| 710 | regards to (4), we fitted the gambin species abundance distribution model, using the 'gambin' |
| 711 | R package and a subsampling procedure to ensure sample sizes were kept consistent (see |
| 712 | Matthews et al. 2014), to the abundances of the species in each assemblage and recorded the |
| 713 | alpha (shape) parameter (Matthews et al. 2014). Finally, to check that spatial species turnover |
| 714 | across the three traps was low, we calculated multisite spatial beta-diversity across the three |
| 715 | assemblages using presence-absence data ('beta.multi' function; Sorensen index) and |
| 716 | abundance data ('beta.multi.abund' function; Bray-Curtis multiple-site dissimilarity), using |
| 717 | the 'betapart' R package (Baselga et al. 2017). As most species were found in all three traps, |
| 718 | it was not possible to use a null model approach that constrained the marginal totals. All |
| 719 | analyses were undertaken using R (version 3.4.3, R Core Team 2017). |

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Plot locations

27°21W 27°16W 27°15W 27°12W 27°9W 27

Figure S1. A map of the ten sampled 50m x 50m plots and areas of native forest on the island of Terceira, Azores.



Figure S2. Example SLAM trap in a plot within a native forest fragment on Terceira Island (Terra-Brava T18), Azores

Table S1. Sampling dates for the nine plots across the five years. A '1' indicates that sampling was undertaken in this plot during this time period, whilst a '0' indicates that sampling was not undertaken during this time period due to either logistical issues or problems with the traps. Codes as in Figure S1: 1- TER-NFBF-T-01; 2- TER-NFBF-T-02; 3- TER-NFBF-TP41; 4- TER-NFPG-T-33; 5- TER-NFSB-T-07; 6- TER-NFSB-T164; 7-TER-NFSB-TE48; 8- TER-NFSB-TE49; 9- TER-NFTB-T-15,

| Plot | 2012 | 2012 | 2013 | 2013 | 2013 | 2014 | 2014 | 2014 | 2014 |
|--------|-------|------|------|------|-------|-------|------|------|------|
| 1101 | Sep. | Dec. | May | Sep. | Dec. | March | June | Sep. | Dec. |
| 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| 2 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| 3 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| 4 | 1 | 0 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| 5 | 1 | 0 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| 6 | 1 | 0 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| 7 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| 8 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| 9 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| D1 - 4 | 2015 | 2015 | 2015 | 2015 | 2016 | 2016 | 2016 | 2016 | |
| Plot | March | Sep. | June | Dec. | March | June | Sep. | Dec. | |
| 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | |
| 2 | 1 | 1 | 1 | 0 | 1 | 1 | 1 | 1 | |
| 3 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | |
| 4 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | |
| 5 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 0 | |
| 6 | 1 | 0 | 1 | 1 | 1 | 1 | 1 | 1 | |
| 7 | 1 | 1 | 1 | 0 | 1 | 1 | 1 | 1 | |
| 8 | 1 | 1 | 1 | 1 | 1 | 0 | 1 | 1 | |
| 9 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | |

Appendix S2 Temporal beta-diversity results using the Carvalho et al. partition framework

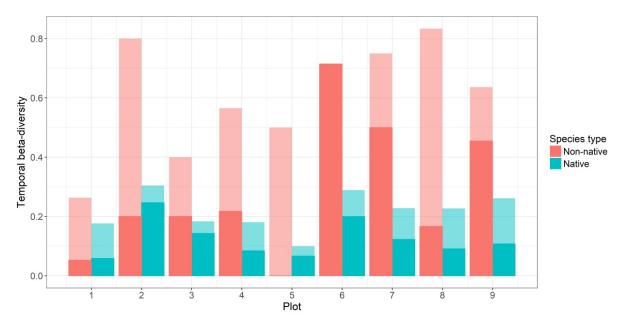


Figure S3. Pairwise temporal beta-diversity values for native and non-native arthropod species across nine plots of native forest on Terceira Island, in the Azores (Plot notation as in Table S1 and Figure S1). For each plot, temporal beta-diversity was calculated using the pooled 2013 samples and the pooled 2016 samples. For each plot, temporal beta-diversity was calculated separately for native species (blue bars) and non-native species (red bars). The height of each bar corresponds to the overall temporal beta-diversity. Overall temporal beta was also partitioned into richness differences and turnover components using the Carvalho partition approach (Carvalho et al. 2012). Within each bar, the dark shaded area rising from the x-axis corresponds to the richness differences component of overall temporal beta diversity. Thus, the lighter shaded area within each bar that rises from the dark shaded area corresponds to the turnover component.

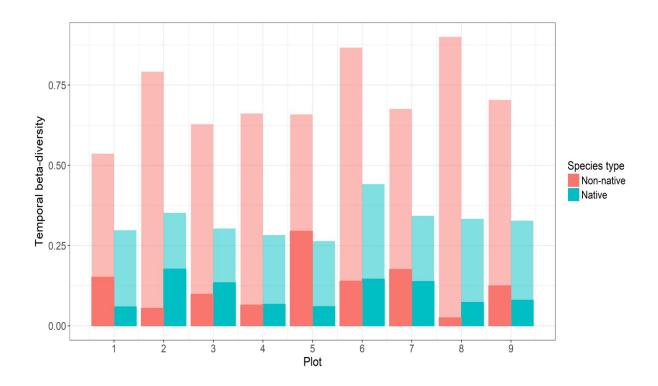


Figure S4. Multiple-time dissimilarity values for native and non-native arthropod species across nine plots of native forest on Terceira Island, in the Azores. For each plot, multiple-time dissimilarity was calculated across four years of data (2013 – 2016). For each plot, temporal beta-diversity was calculated separately for native species (blue bars) and non-native species (red bars). The height of each bar corresponds to the overall multiple-time dissimilarity was. Overall multiple-time dissimilarity was also partitioned into nestedness and turnover components using the Baselga partition approach (Baselga et al. 2017). Within each bar, the dark shaded area rising from the x-axis corresponds to the nestedness component of overall temporal beta diversity. Thus, the lighter shaded area within each bar that rises from the dark shaded area corresponds to the turnover component.

Appendix S4 Sensitivity Analyses Full Results

Table S2. Sampling coverage estimates for the nine plots across the five years. SC is the sampling coverage for the observed data, and the LCI and UCI are the 95% lower and upper confidence intervals, respectively.

| SC | LCI | UCI | Plot | Year | SC | LCI | UCI | Plot | Year |
|-------|-------|-------|------|------|-------|-------|-------|------|------|
| 0.992 | 0.987 | 0.996 | 1 | 2012 | 0.983 | 0.967 | 1.000 | 8 | 2012 |
| 0.986 | 0.981 | 0.991 | 1 | 2013 | 0.987 | 0.981 | 0.993 | 8 | 2013 |
| 0.992 | 0.987 | 0.997 | 1 | 2014 | 0.989 | 0.983 | 0.995 | 8 | 2014 |
| 0.987 | 0.982 | 0.993 | 1 | 2015 | 0.981 | 0.968 | 0.993 | 8 | 2015 |
| 0.986 | 0.981 | 0.992 | 1 | 2016 | 0.981 | 0.970 | 0.991 | 8 | 2016 |
| 0.942 | 0.911 | 0.974 | 2 | 2012 | 0.979 | 0.965 | 0.992 | 9 | 2012 |
| 0.984 | 0.977 | 0.992 | 2 | 2013 | 0.990 | 0.986 | 0.995 | 9 | 2013 |
| 0.985 | 0.975 | 0.995 | 2 | 2014 | 0.985 | 0.976 | 0.993 | 9 | 2014 |
| 0.967 | 0.952 | 0.981 | 2 | 2015 | 0.976 | 0.964 | 0.989 | 9 | 2015 |
| 0.989 | 0.983 | 0.995 | 2 | 2016 | 0.973 | 0.958 | 0.987 | 9 | 2016 |
| 0.970 | 0.956 | 0.985 | 3 | 2012 | | | | | |
| 0.990 | 0.982 | 0.997 | 3 | 2013 | | | | | |
| 0.996 | 0.993 | 1.000 | 3 | 2014 | | | | | |
| 0.972 | 0.959 | 0.985 | 3 | 2015 | | | | | |
| 0.985 | 0.974 | 0.996 | 3 | 2016 | | | | | |
| 0.918 | 0.878 | 0.957 | 4 | 2012 | | | | | |
| 0.991 | 0.987 | 0.995 | 4 | 2013 | | | | | |
| 0.991 | 0.988 | 0.994 | 4 | 2014 | | | | | |
| 0.979 | 0.972 | 0.986 | 4 | 2015 | | | | | |
| 0.985 | 0.978 | 0.992 | 4 | 2016 | | | | | |
| 0.810 | 0.666 | 0.954 | 5 | 2012 | | | | | |
| 0.994 | 0.989 | 0.998 | 5 | 2013 | | | | | |
| 0.989 | 0.982 | 0.996 | 5 | 2014 | | | | | |
| 0.984 | 0.976 | 0.992 | 5 | 2015 | | | | | |
| 0.988 | 0.978 | 0.999 | 5 | 2016 | | | | | |
| 0.902 | 0.845 | 0.958 | 6 | 2012 | | | | | |
| 0.954 | 0.931 | 0.977 | 6 | 2013 | | | | | |
| 0.968 | 0.956 | 0.980 | 6 | 2014 | | | | | |
| 0.965 | 0.947 | 0.984 | 6 | 2015 | | | | | |
| 0.966 | 0.949 | 0.984 | 6 | 2016 | | | | | |
| 0.983 | 0.957 | 1.009 | 7 | 2012 | | | | | |
| 0.977 | 0.965 | 0.990 | 7 | 2013 | | | | | |
| 0.985 | 0.977 | 0.993 | 7 | 2014 | | | | | |
| 0.963 | 0.943 | 0.984 | 7 | 2015 | | | | | |
| 0.910 | 0.867 | 0.953 | 7 | 2016 | | | | | |

Using three SLAM traps in a single plot Using three SLAM traps in a plot rather than one did not result in substantially different sampled communities. Species richness across the three traps in T18 only ranged 49 – 51 species, whilst abundance varied from 1380 to 1739 individuals per trap. The alpha parameter of the gambin distribution was relatively constant between the three traps (2.47, 2.44 and 3.69), indicating similar SAD shapes. 58% of the total sampled species (n = 38) were present in all three traps and overall spatial beta-diversity was low based on both presence/absence data (Sorensen dissimilarity = 0.24) and abundance data (Bray-Curtis multiple-site dissimilarity = 0.31). Thus, we are confident that our general sampling protocol is sufficient to obtain representative samples of the arthropod fauna in each plot.